

$$1) f_X(x) = \begin{cases} \alpha x^2 & , 0 \leq x \leq 10 \\ 0 & , \text{else} \end{cases}$$

a) Find α . By DEFINITION, $\int_{-\infty}^{\infty} f_X(x) dx = 1 = \int_{-\infty}^0 0 \cdot dx + \int_0^{10} \alpha x^2 dx + \int_{10}^{\infty} 0 dx$

$$1 = \int_0^{10} \alpha x^2 dx = \alpha \frac{x^3}{3} \Big|_0^{10} = \alpha \cdot \frac{1000}{3}$$

$$\boxed{\alpha = \frac{3}{1000}}$$

b) Find $P(X > 5)$. This equals $1 - P(X \leq 5)$.

By definition $P(X \leq 5) = F_X(5)$

and $F_X(x) = \int_{-\infty}^x f_X(x) dx$

so $F_X(5) = \int_{-\infty}^0 0 \cdot dx + \int_0^5 \frac{3}{1000} x^2 dx$

$$F_X(5) = \frac{3}{1000} \cdot \frac{x^3}{3} \Big|_0^5 = \frac{3}{1000} \cdot \frac{125}{3} = \frac{125}{1000} = 0.125$$

so $P(X \leq 5) = 0.125$

and $P(X > 5) = 1 - 0.125 = \boxed{0.875}$

c) Find μ_X . $\mu_X = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\infty}^0 0 dx + \int_0^{10} \frac{3}{1000} x^3 dx + \int_{10}^{\infty} 0 dx$

$$= \frac{3}{1000} \cdot \frac{x^4}{4} \Big|_0^{10} = \frac{30,000}{4,000} - 0 = \frac{30}{4} = \boxed{7.5}$$

d) Find σ_X^2 and σ_X . $\sigma_X^2 = E[X^2] - \mu_X^2 = \int_{-\infty}^{\infty} x^2 f_X(x) dx - (\mu_X)^2$

$$= \int_{-\infty}^0 0 dx + \int_0^{10} \frac{3}{1000} x^4 dx + \int_{10}^{\infty} 0 dx - (7.5)^2 = \frac{3}{1000} \cdot \frac{x^5}{5} \Big|_0^{10} - (7.5)^2$$

$$= \frac{3}{1000} \cdot \frac{100,000}{5} - 56.25 = 60 - 56.25 = \boxed{3.75 = \sigma_X^2} \quad \boxed{\sigma_X = 1.94}$$

2) GAUSSIAN DISTRIBUTION OF X WITH $\mu_X = 60$ DAYS

$$\sigma_X = 15 \text{ DAYS}$$

$$a) f_X(x) = \frac{1}{\sigma_X \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x - \mu_X}{\sigma_X} \right)^2} = \boxed{\frac{1}{15\sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x - 60}{15} \right)^2} = f_X(x)}$$

$$b) F_X(x) = \int_{-\infty}^x f_X(x) dx = \frac{1}{15\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{1}{2} \left(\frac{x - 60}{15} \right)^2} dx$$

$$\left[\text{USE THAT } \frac{1}{\sigma_X \sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{(x - \mu_X)^2}{2\sigma_X^2}} dx = \frac{1}{2} \left[1 + \text{erf} \left(\frac{x - \mu_X}{\sigma_X \sqrt{2}} \right) \right] \right]$$

$$\boxed{F_X(x) = \frac{1}{2} \left[1 + \text{erf} \left(\frac{x - 60}{15\sqrt{2}} \right) \right]}$$

$$c) P[40 < X \leq 70] = F_X(70) - F_X(40) = 0.7486 - 0.0918 = 0.6568$$

$$\text{TO GET } F_X(70) = \frac{1}{2} \left[1 + \text{erf} \left(\frac{70 - 60}{15\sqrt{2}} \right) \right] = 0.7486$$

$$F_X(40) = \frac{1}{2} \left[1 + \text{erf} \left(\frac{40 - 60}{15\sqrt{2}} \right) \right] = 0.0918$$

PLUG INTO PROGRAM
SUCH AS MATHECAD,
PROGRAMMABLE
CALCULATOR, ETC.

ALTERNATIVELY, YOU CAN USE STANDARD NORMAL TABLES.

$$\text{where } P(a < X \leq b) = \Phi \left(\frac{b - \mu_X}{\sigma_X} \right) - \Phi \left(\frac{a - \mu_X}{\sigma_X} \right)$$

PLUG THESE INTO
↑
TABLES TO GET Φ VALUES.

$$P(40 < X \leq 70) = \Phi \left(\frac{70 - 60}{15} \right) - \Phi \left(\frac{40 - 60}{15} \right) = \Phi(0.67) - \Phi(-1.33)$$

$$\left[\text{There aren't negatives in the table so use } \Phi(-1.33) = 1 - \Phi(1.33) \right]$$

$$P(40 < X \leq 70) = \Phi(0.67) - [1 - \Phi(1.33)] = 0.7486 - (1 - 0.9082) = 0.6568$$

$$\boxed{P(40 < X \leq 70) = 65.7\%}$$

$$2) d) P[X \geq 30 \text{ days}] = 1 - F_X(30) = 1 - 0.0228 = 0.9772$$

ALTERNATIVELY, USING TABLES

$$P(30 < X \leq \infty) = \Phi(\infty) - \Phi\left(\frac{30-60}{15}\right) = 1 - \Phi(-2.00) \\ = 1 - [1 - \Phi(2.00)] = \Phi(2.00) = 0.9772$$

$$P(30 < X) = 0.9772$$

3) POISSON DISTRIBUTION

$$\lambda = 5 \frac{\text{RAINSTORMS}}{\text{YEAR}}$$

$$P[X=x] = \frac{(\lambda t)^x}{x!} e^{-\lambda t} = \frac{(5t)^x}{x!} e^{-5t}$$

$$a) P[X=0] = \frac{(5 \cdot 1)^0}{0!} e^{-5 \cdot 1} = e^{-5} = 0.00674$$

Next year
 $\therefore t=1 \text{ year}$

$$P[0 \text{ storms next year}] = 0.7\%$$

$$b) P[X=5] = \frac{(5 \cdot 1)^5}{5!} e^{-5 \cdot 1} = \frac{5^5}{5!} e^{-5} = 0.17547$$

Next year
 $\therefore t=1 \text{ year}$

$$P[5 \text{ storms next year}] = 17.5\%$$

$$c) P[X \geq 3] = 1 - P[X=0] - P[X=1] - P[X=2]$$

Next year
 $\therefore t=1 \text{ year}$

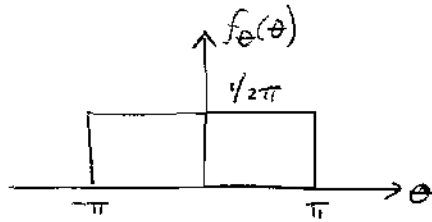
$$P[X=1] = \frac{(5 \cdot 1)^1}{1!} e^{-5 \cdot 1} = 0.03369$$

$$P[X=2] = \frac{(5 \cdot 1)^2}{2!} e^{-5 \cdot 1} = 0.08422$$

$$P[X \geq 3] = 1 - 0.00674 - 0.03369 - 0.08422 = 0.87535$$

$$P[X \geq 3 \text{ STORMS NEXT YEAR}] = 87.5\%$$

$$4) \quad \eta(x, t) = A \sin(\omega_0 t + k_0 x) = A \sin(\omega_0 t + \theta) = \eta(t, \theta)$$



a) ENSEMBLE AVERAGE IS CALCULATED AT A CONSTANT TIME, SO LET $t = t_0$

$$\mu_\theta(t_0) = E\{\eta(t_0, \theta)\} = \int_{-\infty}^{\infty} \eta(t_0, \theta) \cdot f_\theta(\theta) d\theta$$

$$\left[\text{USE EQN (32), PG 5 OF R.V. LECTURE: } \uparrow E[h(x)] = \mu_{h(x)} = \int_{-\infty}^{\infty} h(x) f_x(x) dx \right]$$

$$\mu_\theta(t_0) = \int_{-\infty}^{-\pi} \eta(t_0, \theta) \cdot 0 \cdot d\theta + \int_{-\pi}^{\pi} \eta(t_0, \theta) \cdot \frac{1}{2\pi} \cdot d\theta + \int_{\pi}^{\infty} \eta(t_0, \theta) \cdot 0 \cdot d\theta$$

$$= \int_{-\pi}^{\pi} \eta(t_0, \theta) \frac{1}{2\pi} d\theta = \frac{1}{2\pi} \int_{-\pi}^{\pi} A \sin(\omega_0 t_0 + \theta) d\theta$$

$$= \frac{A}{2\pi} \underbrace{\int_{-\pi}^{\pi} \sin(\omega_0 t_0 + \theta) d\theta}_0 = \boxed{0 = \text{ENSEMBLE AVERAGE}}$$

ENSEMBLE VARIANCE, AGAIN LET $t = t_0$.

$$\text{VAR}(t_0) = E\{[\eta(t_0, \theta) - \mu_\theta(t_0)]^2\} = E\{\eta^2(t_0, \theta)\}$$

$$= \int_{-\infty}^{\infty} \eta^2(t_0, \theta) f_\theta(\theta) d\theta = \int_{-\pi}^{\pi} \frac{1}{2\pi} \cdot A^2 \sin^2(\omega_0 t_0 + \theta) d\theta$$

$$= \frac{A^2}{2\pi} \underbrace{\int_{-\pi}^{\pi} \sin^2(\omega_0 t_0 + \theta) d\theta}_{\pi} = \frac{A^2}{2\pi} \cdot \pi = \boxed{\frac{A^2}{2} = \text{ENSEMBLE VARIANCE}}$$

CORRELATION, LET $t = t_0$ AND τ BE A CONSTANT TIME LAG.

$$R_{\theta\theta}(t_0, t_0 + \tau) = E\{\eta(t_0, \theta) \cdot \eta(t_0 + \tau, \theta)\} = \int_{-\infty}^{\infty} \eta(t_0, \theta) \eta(t_0 + \tau, \theta) f_\theta(\theta) d\theta$$

$$= \int_{-\pi}^{\pi} A \sin(\omega_0 t_0 + \theta) A \sin(\omega_0(t_0 + \tau) + \theta) \frac{1}{2\pi} d\theta$$

$$4a \text{ cont.}) \quad R_{\eta\eta}(t_0, t_0 + \tau) = \frac{A^2}{2\pi} \int_{-\pi}^{\pi} \sin(\omega_0 t_0 + \theta) \sin(\omega_0(t_0 + \tau) + \theta) d\theta$$

$$R_{\eta\eta} = \frac{A^2}{2\pi} \cdot \pi \cos(\omega_0 \tau) = \frac{A^2}{2} \cos(\omega_0 \tau)$$

$$\boxed{R_{\eta\eta} = \frac{A^2}{2} \cos(\omega_0 \tau)}$$

b) TEMPORAL AVERAGE:

$$M^t \{ \eta_i(t) \} = \bar{\eta}^t = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T \eta_i(t) dt$$

$$= \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T A \sin(\omega_0 t + \theta_i) dt = \lim_{T \rightarrow \infty} \frac{1}{T} \cdot \frac{-A}{\omega_0} [\cos(\omega_0 t + \theta_i)]_0^T$$

$$= \lim_{T \rightarrow \infty} \frac{-A}{T \omega_0} [\cos(\omega_0 T + \theta_i) - \cos(\theta_i)] = \boxed{0 = \bar{\eta}^t}$$

TEMPORAL VARIANCE

$$V^t \{ \eta_i(t) \} = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T [\eta_i(t) - \underbrace{M^t \{ \eta_i(t) \}}_0]^2 dt = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T \eta_i^2(t) dt$$

$$= \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T A^2 \sin^2(\omega_0 t + \theta_i) dt = \lim_{T \rightarrow \infty} \frac{A^2}{T} \underbrace{\int_0^T \sin^2(\omega_0 t + \theta_i) dt}_{T/2}$$

$$= \lim_{T \rightarrow \infty} \frac{A^2 T}{T/2} = \lim_{T \rightarrow \infty} \frac{A^2}{2} = \boxed{\frac{A^2}{2} = V^t \{ \eta_i(t) \}}$$

CORRELATION

$$R^t \{ \eta_i(t) \} = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T [\eta_i(t) - \underbrace{M^t \{ \eta_i(t) \}}_0] [\eta_i(t+\tau) - \underbrace{M^t \{ \eta_i(t+\tau) \}}_0] dt$$

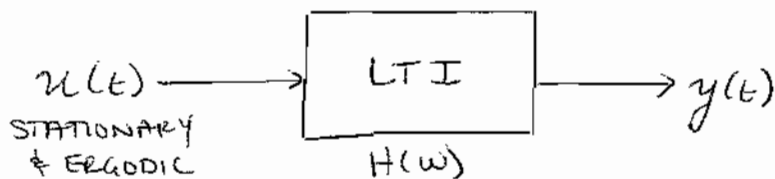
$$= \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T \eta_i(t) \eta_i(t+\tau) dt = \lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T A^2 \sin(\omega_0 t + \theta_i) \sin(\omega_0(t+\tau) + \theta_i) dt$$

$$= \lim_{T \rightarrow \infty} \frac{A^2}{T} \int_0^T \sin(\omega_0 t + \theta_i) \sin(\omega_0(t+\tau) + \theta_i) dt$$

$$\boxed{R^t \{ \eta_i(t) \} = \frac{A^2}{2} \cos(\omega_0 \tau)}$$

∴ ENSEMBLE STATISTICS EQUAL TEMPORAL STATISTICS.

5)



- a) Yes. IF the input to an LTI system is stationary and ergodic, the output is as well.
- b) IF a random process has ensemble statistics that do not depend on time, it is stationary.
- c) Yes. You can have a stationary, non-ergodic process.
No. You cannot have an ergodic, non-stationary process.
- d) Given $S_x(\omega)$ and $H(\omega)$, $S_y(\omega) = |H(\omega)|^2 S_x(\omega)$.
- e) Given $S_x(\omega)$ and $S_y(\omega)$, $|H(\omega)| = \sqrt{\frac{S_y(\omega)}{S_x(\omega)}}$.
- You can find the magnitude of $H(\omega)$, but you would need more information to find the phase of $H(\omega)$.

6) $\sigma = M_0 = 1.25 \text{ m}$ $\bar{T} = 12.4 \text{ sec}$ PLATFORM HEIGHT, $h = A = 3 \text{ m}$

a) $\bar{\eta}(A) = \frac{1}{\bar{T}} \cdot e^{-A^2/2M_0}$

$$\bar{\eta}(3 \text{ m}) = \frac{1}{12.4 \text{ sec}} \cdot e^{\frac{-(3 \text{ m})^2}{2 \cdot (1.25 \text{ m})^2}} = 0.00453 \text{ s}^{-1}$$

$$\begin{aligned} \bar{\eta}(3 \text{ m}) &= 0.00453 \text{ upcrossings per second} \\ &= 0.2716 \text{ upcrossings per minute} \\ &= 16.3 \text{ upcrossings per hour} \end{aligned}$$

b) $\bar{\eta}(h) = \frac{1}{12.4 \text{ sec}} \cdot e^{-h^2/2(1.25 \text{ m})^2} = \frac{1}{1 \text{ hour}} = \frac{1}{60 \text{ min}} = \frac{1}{3600 \text{ sec}}$

$$e^{\frac{-h^2}{3.125 \text{ m}^2}} = \frac{12.4 \text{ sec}}{3600 \text{ sec}} = 0.003444$$

$$\frac{-h^2}{3.125 \text{ m}^2} = \ln [0.003444] = -5.67099$$

$$h^2 = (3.125 \text{ m}^2)(5.67099)$$

$$h^2 = 17.722 \text{ m}^2$$

$$h = 4.21 \text{ m}$$

c) $\bar{\eta}(h) = \frac{1}{12.4 \text{ sec}} e^{-h^2/3.125 \text{ m}^2} = \frac{1}{1 \text{ day}} = \frac{1}{24 \text{ hr}} = \frac{1}{86400 \text{ sec}}$

$$e^{-h^2/3.125 \text{ m}^2} = \frac{12.4 \text{ sec}}{86400 \text{ sec}} = 0.000144$$

$$-h^2/3.125 \text{ m}^2 = \ln [0.000144] = -8.849$$

$$h^2 = (3.125 \text{ m}^2)(8.849) = 27.6533 \text{ m}^2$$

$$h = 5.26 \text{ m}$$