

14.381

PS#6 SOLUTIONS

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1. C&B 5.3

Note that $Y_i \sim \text{Bernoulli}$ with $p_i = P(X_i \geq \mu) = 1 - F(u)$ for each i . Since the Y_i 's are iid Bernoulli,

$$\sum_{i=1}^n Y_i \sim \text{binomial}(n, p) \quad \text{with } p = 1 - F(u)$$

2. C&B 5.8

$$\begin{aligned} (a) \quad & \frac{1}{2n(n-1)} \sum_{i=1}^n \sum_{j=1}^n (x_i - x_j)^2 = \frac{1}{2n(n-1)} \sum_{i=1}^n \sum_{j=1}^n (x_i - \bar{x} + \bar{x} - x_j)^2 \\ &= \frac{1}{2n(n-1)} \sum_{i=1}^n \sum_{j=1}^n [(x_i - \bar{x})^2 - 2(x_i - \bar{x})(x_j - \bar{x}) + (x_j - \bar{x})^2] \\ &= \frac{1}{2n(n-1)} \left\{ \sum_{i=1}^n n(x_i - \bar{x})^2 - 2 \sum_{i=1}^n (x_i - \bar{x}) \underbrace{\sum_{j=1}^n (x_j - \bar{x})}_{=0} + n \sum_{j=1}^n (x_j - \bar{x})^2 \right\} \\ &= \frac{n}{2n(n-1)} \sum_{i=1}^n (x_i - \bar{x})^2 + \frac{n}{2n(n-1)} \sum_{j=1}^n (x_j - \bar{x})^2 \\ &= \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 = S^2 \end{aligned}$$

both sums
are
exactly
the same

(b) It seems induction is the easiest way to proceed, and to simplify things assume, without loss of generality, $\theta_1 = 0$, so $E X_i = 0$.

(i) Prove the equation for $n=4$.

$$S^2 = \frac{1}{24} \sum_{i=1}^4 \sum_{j=1}^4 (x_i - x_j)^2$$

2-5

Need to compute $E[M^2]$ and $E[M]^2$, where M is defined as

$$M = \sum_{i=1}^n \sum_{j=1}^n (x_i - x_j)^2$$

Now $E(M) = 24\theta_2$, because

$$\sum_{j=1}^n (x_i - x_j)^2 = \sum_{j \neq i} (x_i - x_j)^2 \quad \text{which only have 3 terms when } n=4.$$

$$\Rightarrow E(M) = \sum_{i=1}^n \sum_{j=1}^n E[(x_i - x_j)^2] \quad E[(x_i - x_j)^2] = 2\theta_2 \quad i \neq j$$

$$\Rightarrow n \cdot (n-1) \cdot 2\theta_2 = 24\theta_2 \quad \text{if } n=4.$$

Now we need $E(M^2)$: this expected value contains $256 (=4^4)$ terms; we can reduce the computational requirements by looking at terms that have the same structure:

- 112 terms are zero, whenever $i=j$;
- 96 are of the form $E[(x_i - x_j)^2 (x_i - x_k)^2] = \theta_4 + 3\theta_2^2$
- 24 are of the form $E[(x_i - x_j)^4] = 2(\theta_4 + 3\theta_2^2)$
- 24 are of the form $E[(x_i - x_j)^2 (x_k - x_l)^2] = 4\theta_2^2$

We can now compute

$$\text{Var}(S^2) = \frac{1}{24^2} [24 \cdot 2(\theta_4 + 3\theta_2^2) + 96(\theta_4 + 3\theta_2^2) + 24 \cdot 4\theta_2^2 - (24\theta_2)^2]$$

$$= \frac{1}{4} \left[\theta_4 - \frac{1}{3}\theta_2^2 \right]$$

Assume the formula holds for n , and establish it for $n+1$.

Define S_n^2 as S^2 computed using n observations.

We can re write

$$S_{n+1}^2 = \frac{1}{2n(n+1)} \left[\underbrace{\sum_{i=1}^n \sum_{j=1}^n (x_i - x_j)^2}_M + 2 \underbrace{\sum_{k=1}^n (x_k - x_{n+1})^2}_Q \right]$$

then we can write

$$S_{n+1}^2 = \frac{1}{2n(n+1)} (M + 2Q)$$

$$\text{Var}(M) = 4 \cdot n \cdot (n-1)^2 \left[\theta_4 - \frac{n-3}{n-1} \theta_2^2 \right]$$

by induction hypothesis

$$\text{Var}(Q) = n(n+1) \theta_4 - n(n-3) \theta_2^2$$

x_k and x_{n+1} are independent

$$\text{Cov}(M, Q) = 2n(n-1) [\theta_4 - \theta_2^2]$$

Hence,

$$\begin{aligned} \text{Var}(S_{n+1}^2) &= \frac{1}{4n^2(n+1)^2} [\text{Var}(M) + 4\text{Var}(Q) + 4\text{Cov}(M, Q)] \\ &= \frac{1}{4n^2(n+1)^2} \left[4n(n-1)^2 \left(\theta_4 - \frac{n-3}{n-1} \theta_2^2 \right) + 4n(n+1)\theta_4 - 4n(n-3)\theta_2^2 + 8n(n-1)(\theta_4 - \theta_2^2) \right] \\ &= \frac{1}{4n^2(n+1)^2} \left[\theta_4 (4n(n-1)^2 + 4n(n+1) + 8n(n-1)) - \theta_2^2 (4n(n-1)(n-3) + 4n(n-3) + 8n(n-1)) \right] \\ &= \frac{1}{4n^2(n+1)^2} (4n^2(n+1) \theta_4 - 4n(n+1)(n-2) \theta_2^2) \\ &= \frac{1}{n+1} \left(\theta_4 - \frac{(n-2)}{n} \theta_2^2 \right) \end{aligned}$$



4-5

3.- C&B 5.13

Take the hint given in the book and start working from there in the following way:

$$\begin{aligned}
 E(c\sqrt{S^2}) &= c \sqrt{\frac{\sigma^2}{n-1}} \cdot E\left(\sqrt{\frac{S^2(n-1)}{\sigma^2}}\right) = c \sqrt{\frac{\sigma^2}{n-1}} \cdot E(\sqrt{q}) \\
 \text{now remember that } \frac{S^2(n-1)}{\sigma^2} &\sim \chi_{(n-1)}^2 \quad q \sim \chi_{(n-1)}^2 \\
 &= c \sqrt{\frac{\sigma^2}{n-1}} \int_0^\infty \sqrt{q} \frac{1}{\Gamma(\frac{n-1}{2}) 2^{\frac{n-1}{2}}} q^{\frac{(n-1)}{2}-1} e^{-q/2} dq \\
 &= c \sqrt{\frac{\sigma^2}{n-1}} \frac{\Gamma(n/2) \sqrt{2}}{\Gamma((n-1)/2)} \underbrace{\int_0^\infty \frac{1}{\Gamma(n/2) 2^{n/2}} q^{n/2-1} e^{-q/2} dq}_{=1 \text{ pdf of } \chi_n^2} \\
 &= \frac{c \sigma \Gamma(n/2) \sqrt{2}}{\sqrt{n-1} \Gamma((n-1)/2)}
 \end{aligned}$$

$$\begin{aligned}
 c E(\sqrt{S^2}) &= c \cdot \frac{\sigma \Gamma(n/2) \sqrt{2}}{\sqrt{n-1} \Gamma((n-1)/2)} \\
 \Rightarrow c &= \frac{\sqrt{n-1} \Gamma((n-1)/2)}{\Gamma(n/2) \sqrt{2}} //
 \end{aligned}$$

4.- C&B 5.16

$X_i, i=1, 2, 3$ independent $N(\mu, \sigma^2)$

$$(a) \chi_3^2: \sum_{i=1}^3 \left(\frac{X_i - \mu}{\sigma}\right)^2 \sim \chi_3^2$$

$$(b) \frac{\left(\frac{X_1 - \mu}{\sigma}\right)}{\sqrt{\frac{\sum_{i=2}^3 \left(\frac{X_i - \mu}{\sigma}\right)^2}{2}}} \sim t_2$$

Important: the variable in the numerator cannot be in the denominator!

(c) take the answer in (b) and compute its square

$$\frac{\left(\frac{x_1 - 1}{1}\right)^2}{\frac{1}{2} \left\{ \left(\frac{x_2 - 2}{2}\right)^2 + \left(\frac{x_3 - 3}{3}\right)^2 \right\}} \sim \bar{F}_{1,2}$$

5.- C&B 5.23

We will use the conditional of $Z|x$ and the marginal of $X=x$.

$$\begin{aligned} P(Z > z) &= \sum_{x=1}^{\infty} P(Z > z | x) P(X=x) = \sum_{x=1}^{\infty} P(U_1 > z, \dots, U_x > z | x) P(X=x) \\ &= \sum_{x=1}^{\infty} \left\{ \prod_{i=1}^x P(U_i > z) \right\} \cdot P(X=x) \end{aligned}$$

this last step uses the fact that U_i is independent of U_j , $i \neq j$; note that according to the formula X affects the result only through the number of terms that are used.

$$\begin{aligned} &= \sum_{x=1}^{\infty} P(U_i > z)^x P(X=x) = \sum_{x=1}^{\infty} (1-z)^x \frac{1}{(e-1)^x x!} \\ &= \frac{1}{(e-1)} \sum_{x=1}^{\infty} \frac{(1-z)^x}{x!} = \frac{e^{1-z} - 1}{e-1} \quad 0 < z < 1 \end{aligned}$$