

# DISCRETE STOCHASTIC PROCESSES

## Lecture 22

### **Martingales**

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## MARTINGALES

Definition: A **martingale**  $\{Z_n; n \geq 1\}$  is a stochastic process with the properties that  $E[Z_n] < \infty$  for all  $n$  and

$$E[Z_n | Z_{n-1} = z_{n-1}, Z_{n-2} = z_{n-2}, \dots, Z_1 = z_1] = z_{n-1}$$

for all  $n > 1$  and all  $z_1, z_2, \dots, z_{n-1}$ .

The condition  $E[Z_n] < \infty$  is a technical condition that is needed in proofs but not usually critical in applications.

The martingale condition is usually stated as  $E[Z_n | Z_{n-1}, \dots, Z_1] = Z_{n-1}$ .

In general  $E[Z_n | Z_{n-1}, \dots, Z_1]$  is a rv that is a function of  $Z_{n-1}, \dots, Z_1$ , and for a martingale, this function is just  $Z_{n-1}$ .

Specifically, each sample point specifies a sample value  $z_{n-1}$  for  $Z_{n-1}$ , a sample value  $z_{n-2}$  for  $Z_{n-2}$ , etc. and thus each sample point with sample values  $z_{n-1}, \dots, z_1$  specifies a value  $E[Z_n | Z_{n-1} = z_{n-1}, \dots, Z_1 = z_1]$ .

This mapping from sample points to real values is the random variable  $E[Z_n | Z_{n-1}, \dots, Z_1]$ .

**Example 1:** Let  $\{S_n; n \geq 1\}$  be a RW with  $E[X] = 0$ . Then

$$E[S_n | S_{n-1}, \dots, S_1] = E[S_{n-1} + X_n | S_{n-1}, \dots, S_1] = S_{n-1}$$

so a zero mean RW is a martingale (and also has the Markov property).

**Example 1':** Let  $\{S_n; n \geq 1\}$  be a RW with  $E[X] = \bar{X} \neq 0$ . Then  $\{Z_n; n \geq 1\}$ , where  $Z_n = S_n - n\bar{X}$ , is a martingale since

$$\begin{aligned} E[Z_n | Z_{n-1}, \dots, Z_1] &= E[(S_n - n\bar{X}) | Z_{n-1}, \dots, Z_1] = \\ &E[(X_n - \bar{X}) + S_{n-1} - (n-1)\bar{X} | Z_{n-1}, \dots, Z_1] = Z_{n-1} \end{aligned}$$

**Example 2: (Product martingales)** Let  $\{X_i; i \geq 1\}$  be IID with mean 1. Let  $Z_n = X_n X_{n-1} \cdots X_1$ . Then  $\{Z_n; n \geq 1\}$  is a martingale (and again also has the Markov property) because

$$E[Z_n | Z_{n-1}, \dots, Z_1] = E[X_n Z_{n-1} | Z_{n-1}, \dots, Z_1] =$$

$$E[X_n | Z_{n-1}, \dots, Z_1] E[Z_{n-1} | Z_{n-1}, \dots, Z_1] = Z_{n-1}$$

If  $X_n$  is 0 or 2, each with probability 1/2, then  $Z_n$  is  $2^n$  with probability  $2^{-n}$ .  $E[Z_n] = 1$  for all  $n \geq 1$ , but  $\lim_{n \rightarrow \infty} Z_n = 0$  WP1.

Thus:  $1 = \lim_{n \rightarrow \infty} E[Z_n] \neq E[\lim_{n \rightarrow \infty} Z_n] = 0$ .

**Example 2': (Product martingales)** Let  $\{X_i; i \geq 1\}$  be IID with mean  $\bar{X} \neq 0$ .

$$\text{Let } W_n = \frac{X_n X_{n-1} \cdots X_1}{(\bar{X})^n} = \frac{X_n}{\bar{X}} \frac{X_{n-1}}{\bar{X}} \cdots \frac{X_1}{\bar{X}}$$

Then  $\{W_n; n \geq 1\}$  is a martingale.

## Definition: The **Wald martingale**

Let  $\{X_i; i \geq 0\}$  be IID,  $\gamma(r) = \ln E[e^{rX}]$ ,  $S_n = X_1 + \dots + X_n$ . Let  $\{Z_n; n \geq 1\}$  be the random process

$$Z_n = \exp[rS_n - n\gamma(r)]$$

$$(Z_n = \prod_{i=1}^n \exp[rX_i - \gamma(r)] = \prod_{i=1}^n \left( \frac{e^{rX_i}}{E[e^{rX_i}]} \right)).$$

This is a Martingale (and again also has the Markov property) since

$$\begin{aligned} E[Z_n | Z_{n-1}, \dots, Z_1] &= E\left\{ \exp[rS_n - n\gamma(r)] \mid Z_{n-1}, \dots, Z_1 \right\} \\ &= E\left\{ \exp[rX_n - \gamma(r)] \exp[rS_{n-1} - (n-1)\gamma(r)] \mid Z_{n-1}, \dots, Z_1 \right\} \\ &= \left[ g(r) / g(r) \right] E[Z_{n-1} | Z_{n-1}, \dots, Z_1] = Z_{n-1} \end{aligned}$$

**Example 3: (Branching process)** Let the nonnegative, integer-valued iid random variables  $\{Y_{k,n}, k, n \geq 1\}$  represent the number of offspring of individual  $k$  in the  $n$ th generation. Let  $E[Y_{k,n}] = \bar{Y}$ .

Then the number of individuals  $X_{n+1}$  in the  $(n+1)$ st generation is given by

$$X_{n+1} = \sum_{k=1}^{X_n} Y_{k,n}.$$

the initial population is  $X_0 = 1$ , then the sequence of random variables

$$Z_n = \frac{X_n}{(\bar{Y})^n}, \quad n \geq 1$$

is a martingale (and again also has the Markov property), since

$$\begin{aligned} E[Z_{n+1} | Z_n = z_n, \dots, Z_1 = z_1] &= E[Z_{n+1} | Z_n = z_n] = \\ E\left[\frac{X_{n+1}}{(\bar{Y})^{n+1}} \mid Z_n = \frac{X_n}{(\bar{Y})^n} = z_n\right] &= E\left[Z_n \frac{\bar{Y}}{\bar{Y}} \mid Z_n = z_n\right] = z_n \end{aligned}$$

**Lemma:** If  $\{Z_n; n \geq 1\}$  is a martingale, then

$$E[Z_n | Z_i, Z_{i-1}, \dots, Z_1] = Z_i \text{ for all } n > i \geq 1.$$

**Proof:** True by definition of a martingale for  $n = i + 1$ . Take  $n = i + 2$ . Suppose densities exist. Let  $\vec{Z}^i = (Z_i, Z_{i-1}, \dots, Z_1)$ . Then

$$\begin{aligned} E[Z_{i+2} | \vec{Z}^i = \vec{z}^i] &= \int_{z_{i+2}} z_{i+2} f_{Z_{i+2} | \vec{Z}^i}(z_{i+2} | \vec{z}^i) \\ &= \int_{z_{i+2}, z_{i+1}} z_{i+2} f_{Z_{i+2} Z_{i+1} | \vec{Z}^i}(z_{i+2}, z_{i+1} | \vec{z}^i) \\ &= \int_{z_{i+2}, z_{i+1}} z_{i+2} f_{Z_{i+2} | \vec{Z}^{i+1}}(z_{i+2} | \vec{z}^{i+1}) f_{Z_{i+1} | \vec{Z}^i}(z_{i+1} | \vec{z}^i) \\ &= \int_{z_{i+2}} z_{i+1} f_{Z_{i+1} | \vec{Z}^i}(z_{i+1} | \vec{z}^i) = z_i \end{aligned}$$

More compactly,

$$E[Z_{i+2} | \vec{Z}^i] = E_{Z_{i+1}} \left\{ E_{Z_{i+2}} [Z_{i+2} | \vec{Z}^{i+1}] \right\} = E_{Z_{i+1}} [Z_{i+1} | \vec{Z}^i] = Z_i$$

Iterating this proves the lemma. In the same way,  $E[Z_n] = E[Z_1]$  for all  $n \geq 1$ .

## Submartingales and Supermartingales

Definition: A stochastic process  $\{Z_n; n \geq 1\}$  is a **submartingale** if  $E[|Z_n|] < \infty$  for all  $n \geq 1$  and if  $E[Z_n | Z_{n-1}, Z_{n-2}, \dots, Z_1] \geq Z_{n-1}$ .

Definition: A stochastic process  $\{Z_n; n \geq 1\}$  is a **supermartingale** if  $E[|Z_n|] < \infty$  for all  $n \geq 1$  and if  $E[Z_n | Z_{n-1}, Z_{n-2}, \dots, Z_1] \leq Z_{n-1}$ .

Sub and super are strange notation, but that's the convention.

Your winnings in a fair game form a martingale; your winnings at Las Vegas form a supermartingale. By the same argument as before, for any  $n > i \geq 1$ ,

$$E[Z_n | Z_i, Z_{i-1}, \dots, Z_1] \geq Z_i \quad (\text{submartingale})$$

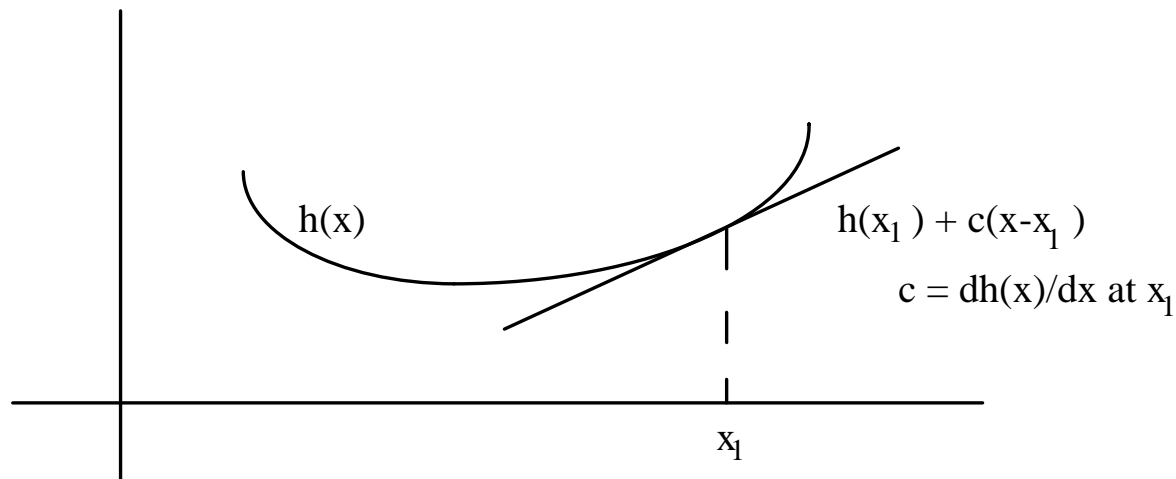
$$E[Z_n | Z_i, Z_{i-1}, \dots, Z_1] \leq Z_i \quad (\text{supermartingale})$$

## Convex Functions and Jensen's Inequality

$h(x)$  is a convex function of the real variable  $x$  if its tangents lie below the function, i.e., if for each  $x_1$ , there is a "slope"  $c$  such that

$$h(x_1) + c(x - x_1) \leq h(x) \text{ for all } x.$$

A function is convex if its second derivative is non-negative, but it can also be convex if the derivatives do not exist everywhere.



**Jensen's Inequality:** If  $h(x)$  is convex and if  $X$  is a rv with  $E[X] < \infty$ , then

$$h(E[X]) \leq E[h(X)].$$

**Proof:** For  $x_1 = E[X]$ , we have  $h(E[X]) + c(x - E[X]) \leq h(x)$  for all  $x$ .

Thus,  $h(E[X]) + c(X - E[X]) \leq h(X)$ .

Taking expected values,  $h(E[X]) \leq E[h(X)]$ .

If  $X$  takes on values  $a$  and  $b$  with equal probability, and  $h$  is a convex function then Jensen says that  $h(a/2 + b/2) \leq h(a)/2 + h(b)/2$ ; i.e., the chord from  $h(a)$  to  $h(b)$  is on top of the curve.

**Theorem:** If  $\{Z_n; n \geq 1\}$  is a martingale and  $h$  is convex, then  $\{h(Z_n); n \geq 1\}$  is a submartingale.

**Proof:**  $E[h(Z_n) | Z_{n-1}, \dots, Z_1] \geq h(E[Z_n | Z_{n-1}, \dots, Z_1]) = h(Z_{n-1})$

**Examples:** If  $\{Z_n; n \geq 1\}$  is a martingale, then  $\{|Z_n|; n \geq 1\}$  is a submartingale,  $\{Z_n^2; n \geq 1\}$  is a submartingale, and  $\{\exp(rZ_n); n \geq 1\}$  is a submartingale.

## Properties

### Martingale

$$E\{Z_n | Z_{n-1}, \dots, Z_1\} = Z_{n-1}$$

### Submartingale

$$E\{Z_n | Z_{n-1}, \dots, Z_1\} \geq Z_{n-1}$$

### Supermartingale

$$E\{Z_n | Z_{n-1}, \dots, Z_1\} \leq Z_{n-1}$$

For  $n > i \geq 1$

$$E\{Z_n | Z_i, \dots, Z_1\} = Z_i$$

$$E\{Z_n | Z_i, \dots, Z_1\} \geq Z_i$$

$$E\{Z_n | Z_i, \dots, Z_1\} \leq Z_i$$

$$E\{Z_n\} = E\{Z_i\}$$

$$E\{Z_n\} \geq E\{Z_i\}$$

$$E\{Z_n\} \leq E\{Z_i\}$$

$$E\{Z_n\} = E\{Z_1\}$$

$$E\{Z_n\} \geq E\{Z_1\}$$

$$E\{Z_n\} \leq E\{Z_1\}$$

## Example: Positive Product Martingales

Let  $\{X_i; i \geq 1\}$  be positive IID random variables with mean 1. Let  $Z_n = X_n X_{n-1} \dots X_1$ . Then  $\{Z_n; n \geq 1\}$  is a martingale. Positive product martingales represent a class of fair games in which a gambler bets 1 dollar, receives  $X_1$  dollars after the first game, bets all his holdings on the second game, and then continues the process. His fortune after  $n$  games is  $Z_n > 0$ .

Inspired by Problem 1.26 in Gallager, we consider the **random walk**

$$Y_n = \ln(Z_n) = \ln X_1 + \ln X_2 + \dots + \ln X_n.$$

Since the logarithm is a **concave** function, it follows from Jensen's inequality that

$$E\{\ln(X)\} \leq \ln(E\{X\}) = \ln(1) = 0,$$

and  $\{Y_i; i \geq 1\}$  is a supermartingale. This inequality is strict if  $\text{Var}(X) > 0$ , and we suppose this is the case. By the strong law,

$$\lim_{n \rightarrow \infty} \left( \frac{Y_n}{n} \right) = E \{ \ln(X) \} < 0, \text{ w.p. } 1,$$

so for **any**  $\beta < 0$ ,

$$\lim_{n \rightarrow \infty} P \{ Y_n < \beta \} = \lim_{n \rightarrow \infty} P \{ \ln(Z_n) < \beta \} = \lim_{n \rightarrow \infty} P \{ Z_n < e^\beta \} = 1,$$

i.e., his fortune eventually shrinks out of sight w.p.1. We've seen an instance of this behavior before in the binary product martingale, i.e., the "double or nothing" nonnegative product martingale where  $P \{ X_i = 0 \} = P \{ X_i = 2 \} = 1/2$ .

While he certainly loses in the long run, the gambler still might make money if he stops when he is ahead, i.e., when his fortune  $Z_n$  equals or exceeds  $e^\alpha > 1$ :

$$P \left\{ \sup_n Y_n \geq \alpha \right\} \leq e^{-r^* \alpha},$$

where

$$g_{\ln(X)}(r) = E \{ e^{r \ln X} \} = E \{ X^r \},$$

so  $r^* = 1$  and

$$P \left\{ \sup_n [Y_n = \ln(Z_n)] \geq \alpha \right\} \leq e^{-\alpha}$$

$$P\left\{\sup_n Z_n \geq e^\alpha\right\} \leq e^{-\alpha},$$

i.e., letting  $e^\alpha = b > 1$ ,

$$P\left\{\sup_n Z_n \geq b \geq 1\right\} \leq 1/b.$$

The likelihood of *ever* doubling your money in any fair game is at most  $1/2$ . This is vastly stronger than the Markov inequality, which states that for each  $k$ ,

$$P\{Z_k \geq b\} \leq E\{Z_k\}/b = 1/b.$$