



MIT Sloan School of Management

MIT Sloan School Working Paper 4785-10

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Principal Components as a Measure of Systemic Risk¹

Mark Kritzman², Yuanzhen Li³, Sebastien Page⁴, and Roberto Rigobon⁵

THIS VERSION: June 28, 2010

The U. S. government's failure to provide adequate oversight and prudent regulation of the financial markets, together with excessive risk taking by some financial institutions, pushed the world financial system to the brink of systemic failure in 2008. As a consequence of this near catastrophe, both regulators and investors have become keenly interested in developing tools for monitoring systemic risk. But this is easier said than done. Securitization, private transacting, complexity, and "flexible" accounting⁶ prevent us from directly observing the many explicit linkages of financial institutions. As an alternative, we introduce a measure of implied systemic risk called the absorption ratio, which equals the fraction of the total variance of a set of asset returns explained or "absorbed" by a fixed number of eigenvectors.⁷ The absorption ratio captures the extent to which markets are unified or tightly coupled. When markets are tightly

¹ We thank Timothy Adler, Robin Greenwood, and participants of seminars at the European Quantitative Forum, the International Monetary Fund, PIMCO, QWAFEFW, and State Street Associates for helpful comments.

² Windham Capital Management, LLC and MIT, Sloan School

³ Windham Capital Management, LLC

⁴ State Street Associates

⁵ MIT, Sloan School and NBER

⁶ Euphemism for accounting practices such as Lehman Brothers' use of swaps to conceal \$50 billion of debt shortly before its demise.

⁷ We could instead calculate the number of eigenvectors required to explain a fixed percentage of variance, but for no particular reason we chose to fix the number of eigenvectors.

coupled, they are more fragile in the sense that negative shocks propagate more quickly and broadly than when markets are loosely linked.

We offer persuasive evidence that the absorption ratio effectively captures market fragility.

We show that:

1. Most significant U.S. stock market drawdowns were preceded by spikes in the absorption ratio.
2. Stock prices, on average, depreciated significantly following spikes in the absorption ratio and, on average, appreciated significantly in the wake of sharp declines in the absorption ratio.
3. The absorption ratio was a leading indicator of the U.S. housing market bubble.
4. The absorption ratio systematically rose in advance of market turbulence.
5. Important milestones throughout the global financial crisis coincided with shifts in the absorption ratio.

We proceed as follows. In Part I we provide a literature review of systemic risk and related topics. In Part II we provide a formal description of the absorption ratio. In Part III we present historical estimates of the absorption ratio for a variety of asset markets, and we show how it relates to asset prices, financial turbulence, the global financial crisis, and financial contagion. We summarize in Part IV and suggest how regulators and investors might use the absorption ratio as an early warning signal of market stress.

Part I: Literature Review

De Bandt and Hartmann (2000) provide an extensive review of the literature on systemic risk. Most studies in this review focus on contagion and “financial fragility,” and the literature on contagion itself is quite rich (see, for example, Pavola and Rigobon, 2008). Recently, the IMF included a chapter on detecting systemic risk in its Global Financial Stability Report (2009), which stated, “The current crisis demonstrates the need for tools to detect systemic risk,” as well as, “Being able to identify systemic events at an early stage enhances policymakers’ ability to take necessary exceptional steps to contain the crisis.” As a simple starting point, the IMF report suggests monitoring conditional (stress) correlations.

In a related study, Billio, Getmansky, Lo, and Pelizzon (2010) show that correlations increase during market crashes. Prior studies have shown that exposure to different country equity markets offers less diversification in down markets than in up markets.⁸ The same is true for global industry returns (Ferreira and Gama, 2004), individual stock returns (Ang, Chen, and Xing, 2002, Ang and Chen, 2002, and Hong, Tu, and Zhou, 2003), hedge fund returns (Van Royen, 2002a), and international bond market returns (Cappiello, Engle, and Sheppard, 2006).

Both the IMF’s Global Financial Stability Report (2009) and Billio, Getmansky, Lo, and Pelizzon (2010) suggest that an important symptom of systemic risk is the presence of sudden regime shifts. Investors have long recognized that economic conditions frequently undergo abrupt changes. The economy typically oscillates between:

-a steady, low volatility state characterized by economic growth; and

⁸ See, for example, Ang and Bekaert, 2002, Kritzman, Lowry, and Van Royen, 2001, and Baele, 2003, on regime shifts; Van Royen, 2002b, and Hyde, Bredin, and Nguyen, 2007 on financial contagion; and Longin and Solnik, 2001, Butler and Joaquin, 2001, Campbell, Koedijk, and Kofman, 2002, Cappiello, Engle, and Sheppard, 2006, and Hyde, Bredin, and Nugyen, 2007 on correlation asymmetries.

-a panic-driven, high volatility state characterized by economic contraction.

Evidence of such regimes has been documented in short-term interest rates (Gray, 1996, Ang and Bekaert, 2002, Smith, 2002), GDP or GNP (Hamilton, 1989, Goodwin, 1993, Luginbuhl and de Vos, 1999, Lam, 2004), inflation (Kim, 1993, Kumar and Okimoto, 2007), and market turbulence (Chow, Jacquier, Kritzman, and Lowry, 1999, Kritzman, Lowry, and Van Royen, 2001, and Kritzman and Li, 2010).

Billio, Getmansky, Lo, and Pelizzon (2010) independently applied principal components analysis to determine the extent to which several financial industries became more unified across two separate regimes. They found that the percentage of the total variance of these industries explained by a single factor increased from 77% during the 1994-2000 period to 83% during the 2001-2008 period. We instead apply principal components analysis to several broad markets and estimate the fraction of total market variance explained by a finite number of factors on a rolling basis throughout history. We call this measure the absorption ratio. We also introduce a standardized measure of shifts in the absorption ratio, and we analyze how these shifts relate to changes in asset prices and financial turbulence. By applying a moving window in our estimation process, we account for potential changes in the risk factors over time. Because Lo, et al. divide history into only two periods, they assume implicitly that these periods are distinct regimes and are stationary within themselves.

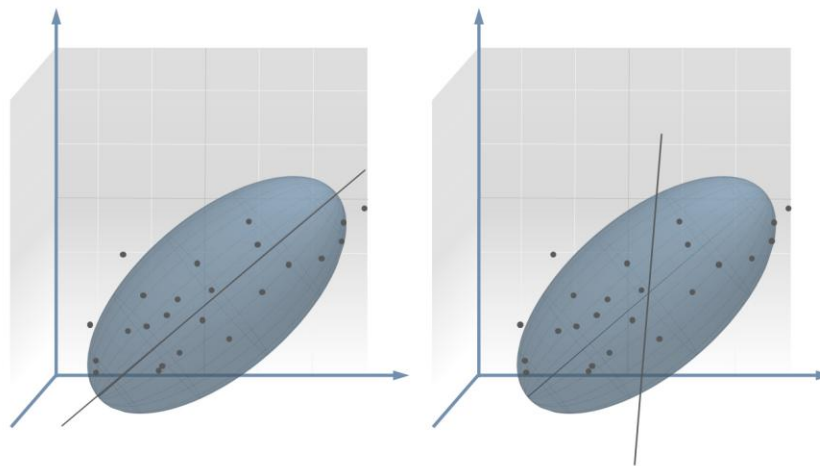
Part II: The Absorption Ratio

Consider a covariance matrix of asset returns estimated over a particular time period. The first eigenvector is a linear combination of asset weights that explains the greatest fraction of the

assets' total variance. The second eigenvector is a linear combination of asset weights orthogonal to the first eigenvector that explains the greatest fraction of leftover asset variance; that is, variance not yet been explained or absorbed by the first eigenvector. The third eigenvector and beyond are identified the same way. They absorb the greatest fraction of leftover variance and are orthogonal to preceding eigenvectors.

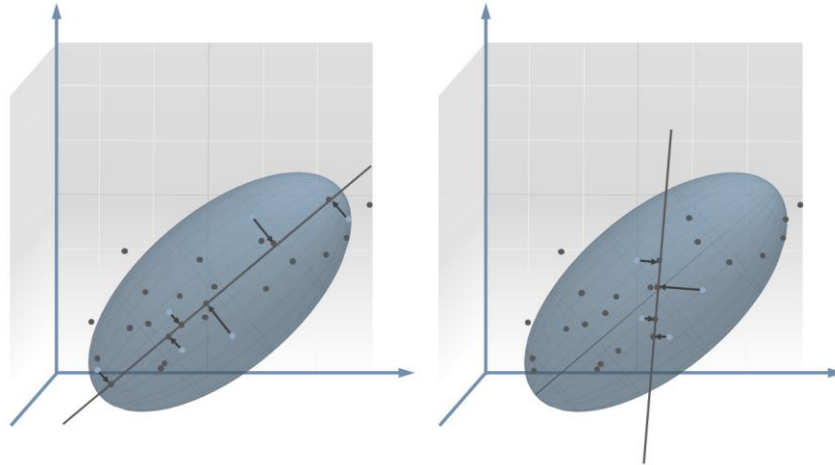
It is perhaps more intuitive to visualize eigenvectors. The left panel of Figure 1 shows a three-dimensional scatter plot of asset returns with a vector piercing the observations. Each observation is the intersection of returns of three assets for a given period, which might be a day, a month, or a year, for example. This vector represents a linear combination of the assets and is a potential eigenvector. The right panel of Exhibit 1 shows the same scatter plot of asset returns but with a different vector piercing the observations.

Exhibit 1: Three Dimensional Scatter Plots of Asset Returns



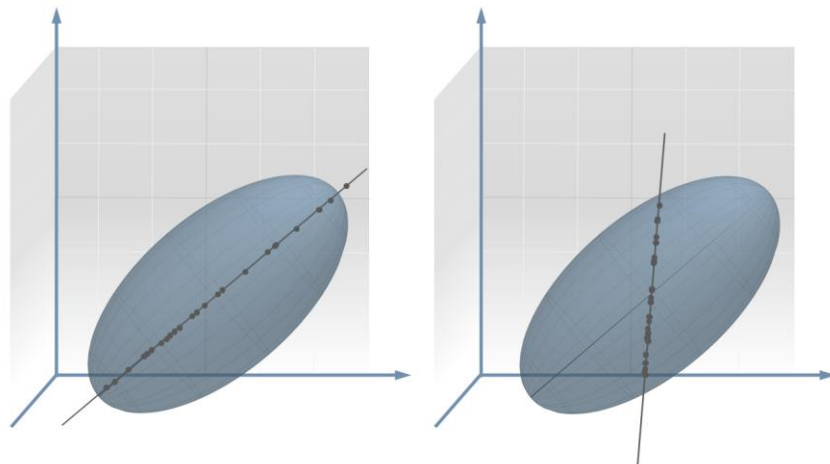
Of all the potential vectors piercing this scatter plot, we determine the first eigenvector by perpendicularly projecting the observations onto each potential eigenvector.

Exhibit 2: Projection of Observations onto Vectors



The first eigenvector is the one with the greatest variance of the projected observations, as shown in Figure 2.⁹

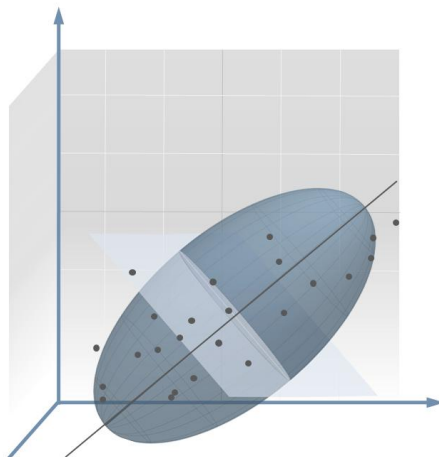
Exhibit 3: First Eigenvector



In order to identify the second eigenvector, we first consider a plane passing through the scatter plot that is orthogonal to the first eigenvector.

⁹There are a variety of techniques to identify eigenvectors. We can use matrix algebra to identify eigenvectors given a small set of observations. For larger data sets, it is more efficient to resort to numerical procedures.

Exhibit 4: Second Eigenvector



The second eigenvector must lie on this orthogonal plane, thereby limiting our search. It is the vector that yields the second highest variance of projected observations. We find the third eigenvector in the same fashion. It is the vector that yields the third greatest variance and is orthogonal to the first two vectors. These three eigenvectors together explain the total variance of the assets.¹⁰

We may or may not be able to associate these eigenvectors with observable economic or financial variables. In some cases the asset weights of the eigenvector may suggest an obvious factor. For example, if we were to observe short exposures to the airline industry and other industries that consume fuel and long exposures to the oil industry and other industries that profit from rising oil prices, we might conclude that this eigenvector is a proxy for the price of oil. Alternatively, an eigenvector may reflect a combination of several influences that came together in a particular way unique to the chosen sample of assets, in which case the factor may not be

¹⁰ The number of eigenvectors never exceeds the number of assets; however, total variation could be explained by fewer eigenvectors than assets to the extent assets are redundant. To be precise, the total number of eigenvectors equals the rank of the covariance matrix.

definable other than as a statistical artifact. Moreover, the composition of eigenvectors may not persist through time. Sources of risk are likely to change from period to period.

In some applications this artificiality or non-stationarity would be problematic. If our intent were to construct portfolios that were sensitive to a particular source of risk, then we would like to be able to identify it and have some confidence of its persistence as an important risk factor. But here our interest is not to interpret sources of risk; rather we seek to measure the extent to which sources of risk are becoming more or less compact.

The particular measure we use as an indicator of systemic risk is the absorption ratio, which we define as the fraction of the total variance of a set of assets explained or absorbed by a finite set of eigenvectors, as shown.

$$AR = \frac{\sum_{i=1}^n \sigma_{E_i}^2}{\sum_{j=1}^N \sigma_{A_j}^2} \quad (1)$$

where,

AR: Absorption ratio

N: number of assets

n: number of eigenvectors used to calculate absorption ratio

$\sigma_{E_i}^2$: variance of the i-th eigenvector, sometimes called eigenportfolio

$\sigma_{A_j}^2$: variance of the j-th asset

A high value for the absorption ratio corresponds to a high level of systemic risk, because it implies the sources of risk are more unified. A low absorption ratio indicates less systemic risk, because it implies the sources of risk are more disparate. We should not expect high systemic risk necessarily to lead to asset depreciation or financial turbulence. It is simply an

indication of market fragility in the sense that a shock is more likely to propagate quickly and broadly when sources of risk are tightly coupled.

The Absorption Ratio versus Average Correlation

One might suspect that the average correlation of the assets used to estimate the absorption ratio provides the same indication of market unity, but it does not. Unlike the absorption ratio, the average correlation fails to account for the relevance of the asset correlations that make up the average.

Exhibit 5: Shift in Correlations and Volatilities

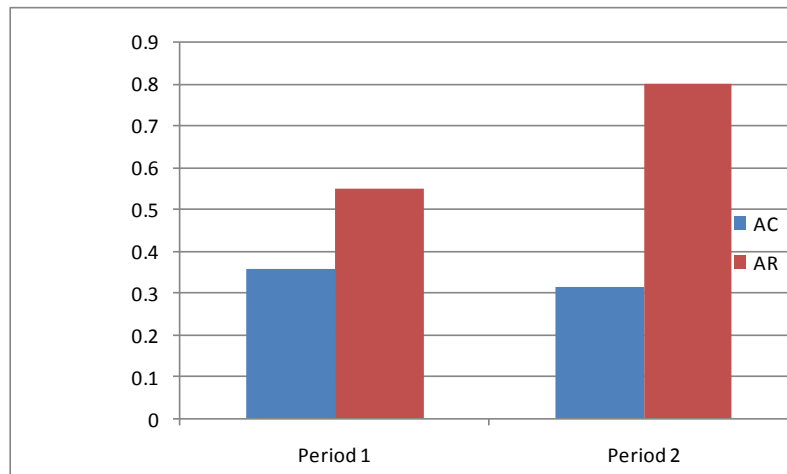
Absorption Ratio versus Average Correlation					
Period 1	Correlations				Standard
Assets	1	2	3	4	Deviations
1	1.00	0.12	-0.01	0.01	35.16%
2	0.12	1.00	-0.04	-0.03	35.07%
3	-0.01	-0.04	1.00	0.82	4.95%
4	0.01	-0.03	0.82	1.00	5.02%

Period 2	Correlations				Standard
Assets	1	2	3	4	Deviations
1	1.00	0.64	-0.05	-0.01	34.46%
2	0.64	1.00	-0.05	-0.03	34.04%
3	-0.05	-0.05	1.00	0.03	4.92%
4	-0.01	-0.03	0.03	1.00	4.88%

Exhibit 5 shows an increase in the correlation of two hypothetical assets with relatively high volatility and a decrease in the correlation of two hypothetical assets with relatively low volatility. It turns out that although the average correlation decreases slightly from one period to the next, the absorption ratio increases sharply, as shown in Exhibit 6. The key distinction is

that the absorption ratio accounts for the relative importance of each asset's contribution to systemic risk whereas the average correlation does not.

Exhibit 6: Shift in Average Correlation and Absorption Ratio



Part III: Empirical Analysis the Absorption Ratio

The Absorption Ratio and Stock Returns

In order to estimate the absorption ratio, we use a window of 500 days to estimate the covariance matrix and eigenvectors, and we fix the number of eigenvectors at approximately 1/5th the number of assets in our sample.^{11 12} The variances, $\sigma_{E_i}^2$ and $\sigma_{A_j}^2$ in Equation (1), are calculated with exponential weighting. This approach assumes that the market's memory of prior events

¹¹ In principle, we should condition the number of eigenvectors on the rank of the covariance. Because the covariance matrices in our analysis are nearly full rank, we are effectively doing this.

¹² As an alternative to measuring market unity by the fraction of total variance explained by a subset of eigenvectors, we could construct a Herfindahl index, by which we square the fraction of variance explained by each of the eigenvectors and sum the squared values. Our experiments show that this latter approach is significantly less informative than our method. We present this evidence in the appendix.

fades away gradually as these events recede further into the past. The half time of the exponential weight decay is set to be half of the window; that is, 250 days.

Exhibit 7 shows a time series of the absorption ratio estimated from the returns of the 51 U.S. industries (hence 10 eigenvectors) in the MSCI USA index based on trailing 500 day overlapping windows, along with the level of MSCI USA price index from January 1, 1998 through January 31, 2010.

Exhibit 7: Absorption Ratio and U.S. Stock Prices

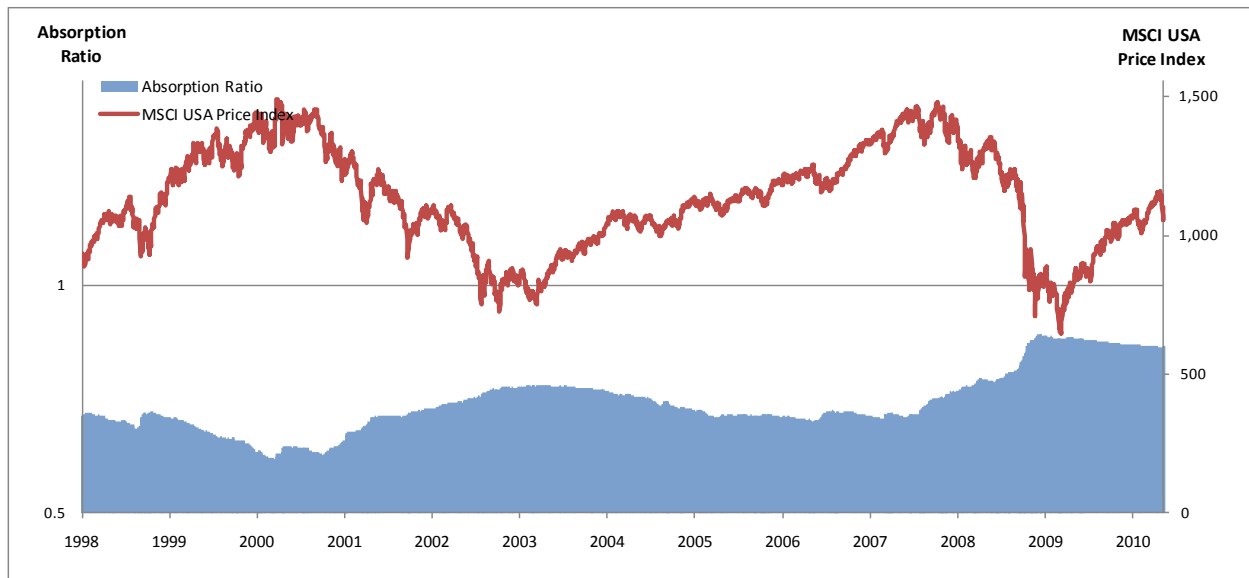


Exhibit 7 shows a distinct inverse association between the level of the absorption ratio and the level of U.S. stock prices. It also reveals that the absorption ratio increased sharply to its highest level ever during the global financial crisis of 2008, coincident with a steep decline in stock prices, and that although stock prices have partially recovered as of the second quarter of 2010, the absorption ratio has fallen only slightly. This continued high level for the absorption

ratio, while perhaps worrisome, does not necessarily foretell a renewed selloff in stocks. It does suggest, however, that the U.S. stock market remains extremely fragile and therefore highly vulnerable to negative shocks.

A casual review of absorption ratio along side stock prices suggests a coincident relationship, which perhaps casts doubt on the notion that the absorption ratio might be useful as a signal of impending trouble. Exhibit 8 sheds some light on this question. It shows the fraction of significant drawdowns preceded by a spike in the absorption ratio. We first compute the moving average of the absorption ratio over 15 days and subtract it from the moving average of the absorption over one year. We then divide this difference by the standard deviation of the one-year absorption ratio, as shown. We call this measure the standardized shift in the absorption ratio.

$$\Delta AR = (AR_{15 \text{ Day}} - AR_{1 \text{ Year}}) / \sigma \quad (2)$$

where,

ΔAR = Standardized shift in absorption ratio

$AR_{15 \text{ Day}}$ = 15-day moving average of absorption ratio

$AR_{1 \text{ Year}}$ = 1-year moving average of absorption ratio

σ = standard deviation of one-year absorption ratio

As Exhibit 8 shows, all of the 1% worst monthly drawdowns were preceded by a one-standard deviation spike in the absorption ratio, and a very high percentage of other significant drawdowns occurred after the absorption ratio spiked.

Exhibit 8: Absorption Ratio and Drawdowns

Fraction of drawdowns preceded by spike in AR			
	1% Worst	2% Worst	5% Worst
1 Day	84.85%	87.69%	70.81%
1 Week	84.85%	83.08%	75.78%
1 Month	100.00%	98.46%	89.44%

1 standard deviation, 15 days / 1 year
1/1/1998 through 5/10/2010

We should not conclude from this exhibit that a spike in the absorption ratio reliably leads to a significant drawdown in stock prices. In many instances, stocks performed well following a spike in the absorption ratio. We would be correct to conclude, though, that a spike in the absorption ratio is a near necessary condition for a significant drawdown, just not a sufficient condition. Again, a high absorption ratio is merely an indication of market fragility.

Even though a spike in the absorption ratio does not always lead to a major drawdown in stock prices, on average stocks perform much worse following spikes in the absorption ratio than they do in the wake of a sharp drop in the absorption ratio. Exhibit 9 shows the average annualized one-day, one-week, and one-month returns following a one-standard deviation increase or decrease in the 15-day absorption ratio relative to the one-year absorption ratio.

Exhibit 9: Absorption Ratio and Subsequent Returns

Annualized return after extreme AR			
	1 Sigma Increase	1 Sigma Decrease	Difference
1 Day	-8.28%	9.27%	-17.56%
1 Week	-8.44%	10.06%	-18.50%
1 Month	-5.86%	12.16%	-18.02%

1 standard deviation, 15 days / 1 year
1/1/1998 through 5/10/2010

Exhibit 9 offers compelling evidence that significant increases in the absorption ratio are followed by significant stock market losses on average, while significant decreases in the absorption ratio are followed by significant gains. This differential performance suggests that it might be profitable to reduce stock exposure subsequent to an increase in the absorption ratio and to raise exposure to stocks after the absorption ratio falls, which is what we next test.

Exhibit 10 shows the performance of a dynamic trading strategy in which the stock exposure of an otherwise equally weighted portfolio of stocks and government bonds is raised to 100% following a one-standard deviation decrease in the 15-day absorption ratio relative to the one-year absorption ratio and reduced to 0% following a one-standard deviation increase. These rules are summarized below.

Absorption Ratio	Stocks/Bonds
$-1\sigma \geq AR \leq +1\sigma$	50/50
$AR > +1\sigma$	0/100
$AR < -1\sigma$	100/0

These rules are applied daily with a one-day lag following the signal for the period January 1, 1998 through January 31, 2010 using the MSCI USA stock index and Treasury bonds.¹³ Exhibit 10 shows that these rules triggered only 1.72 trades per year on average, which should not be surprising given that one-standard deviation events occur infrequently. Nonetheless, these infrequent shifts improved return by more than 4.5% annually while increasing risk by only 0.61%, thereby raising the return/risk ratio from 0.47 to 0.83.

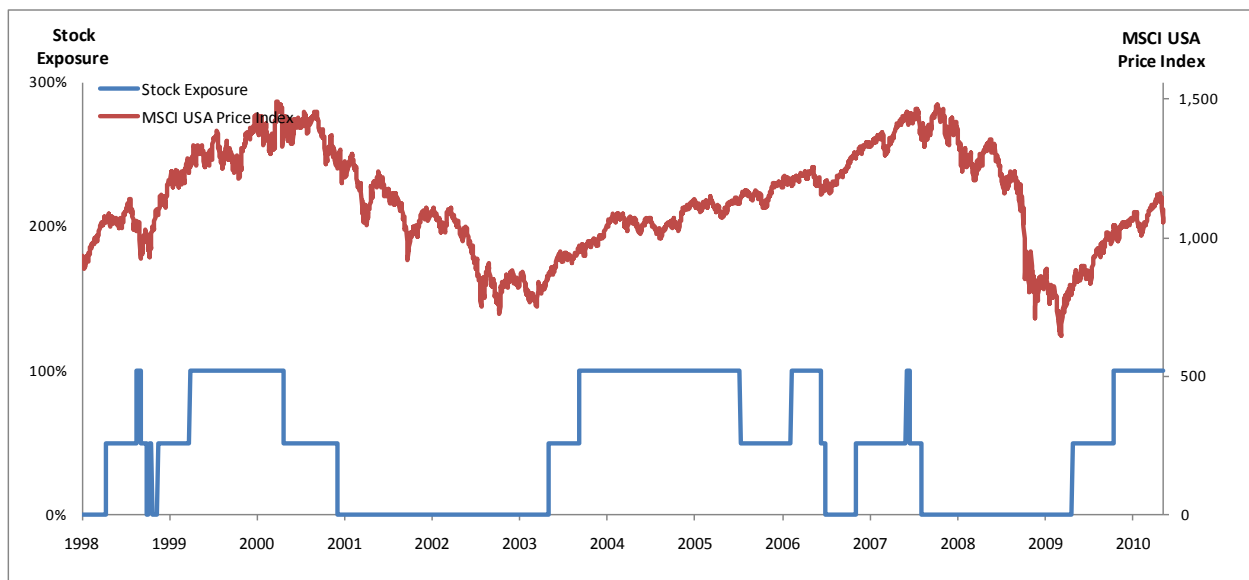
Exhibit 10: Absorption Ratio as a Market Timing Signal

Performance: 100/0 versus 0/100		
	Dynamic	50/50
Return	9.58%	5.08%
Risk	11.50%	10.89%
Return/Risk	0.83	0.47
Turnover		86.01%
Number of Trades		1.72
1/1/1998 through 5/10/2010		

¹³ We require two years to estimate the covariance matrix and eigenvectors and another year to estimate the standard deviation of the one-year moving average; hence our data begins January 1, 1995.

Although most investors might be reluctant to shift entirely in or out of stocks given a single signal, this experiment does offer persuasive evidence of the potential value of the absorption ratio as a market timing signal. Exhibit 11 reveals that the absorption ratio would have kept investors out of stocks during much of the dot com meltdown as well as the global financial crisis. It also reveals that the absorption ratio produced some false positives, but not enough to offset its net beneficial effect.

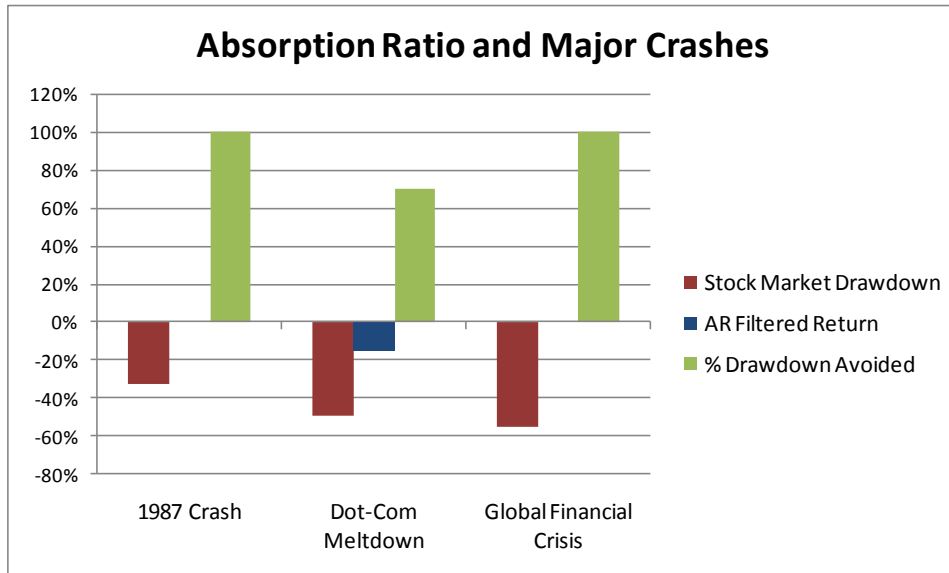
Exhibit 11: Absorption Ratio Stock Exposure



Although daily MSCI industry data extends back only to 1995, Data Stream offers less granular industry data which allows us to include the 1987 stock market crash in our analysis. Exhibit 12 shows how well the absorption ratio served as an early warning signal of the great stock market crashes of the past three decades. If investors withdrew from stocks within one day of a one standard deviation spike in the standardized shift of the absorption ratio, they would

have avoided all of the losses associated with the 1987 crash and the global financial crisis, and nearly 70% of the losses associated with the Dot Com meltdown.

Exhibit 12: The Absorption Ratio as an Early Warning Signal of Crashes



This trading rule appears to improve performance in other the stock markets as well. Exhibits 13 and 14 show that that the absorption ratio estimated from stock returns in Canada, Germany, Japan, and the U.K. and applied as a timing signal in those markets yielded similar improvement in total return as well as risk-adjusted return.¹⁴

¹⁴ As with the U.S. stock market, we set the number of eigenvectors at about 1/5th the number of industries.

Exhibit 13: Global Performance of Absorption Ratio

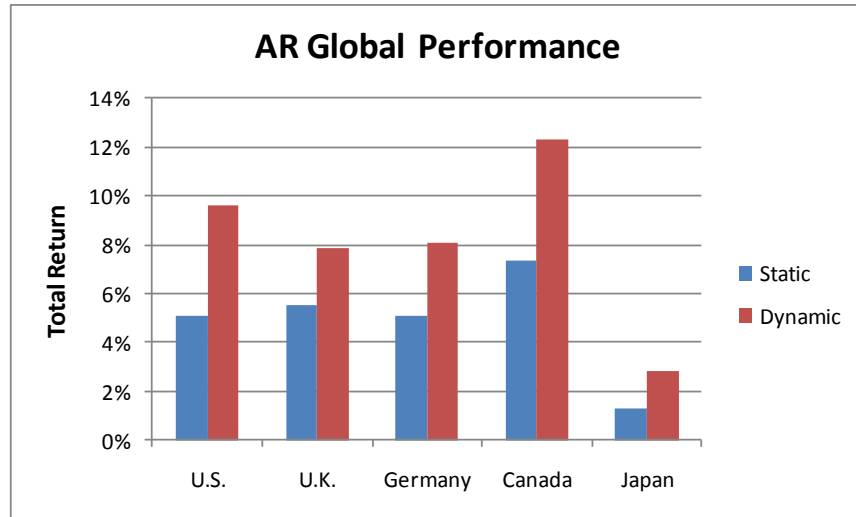
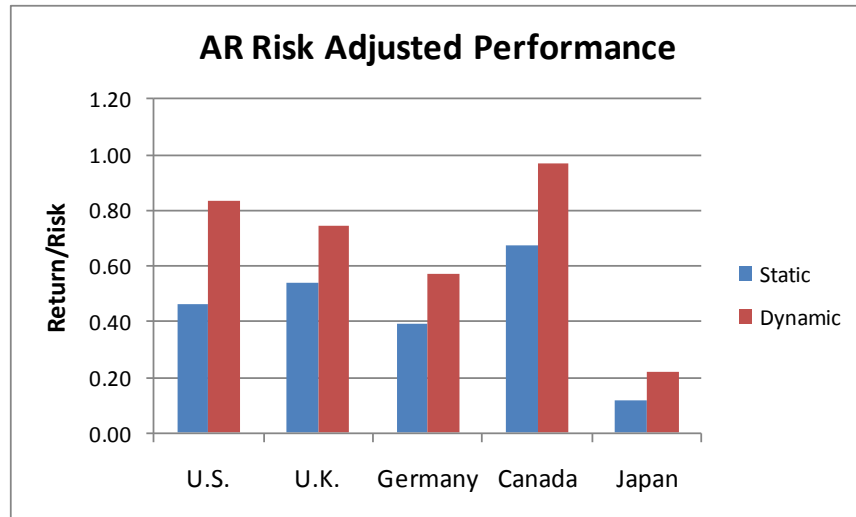


Exhibit 14: Global Performance of Absorption Ratio

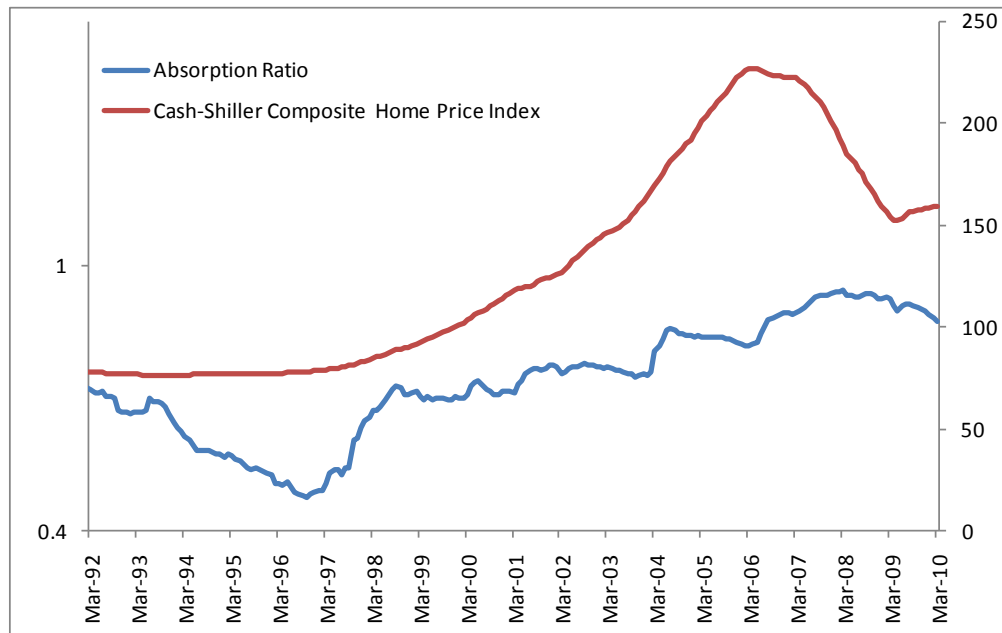


The Absorption Ratio and the Housing Bubble

Former Federal Reserve Chairman Alan Greenspan stated that although regional housing markets often showed signs of unsustainable speculation resulting in local housing bubbles, he did not expect a national U.S. housing bubble.¹⁵ Had the Fed examined the absorption ratio of the U.S. housing market, they might have learned that regional housing markets were becoming more and more tightly coupled as early as 1998, setting the stage for a national housing bubble.

Exhibit 15 shows the absorption ratio estimated from 14 metropolitan housing markets in the United States, along with an index of the Case-Shiller 10-City National Composite Index.¹⁶

Exhibit 15: The Absorption and the National Housing Bubble



¹⁵ He has since rejected this view.

¹⁶ The returns are computed from the Case-Shiller Indexes that go back to 1987. The metropolitan areas include : Los Angeles, San Diego, San Francisco, Denver, Washington, DC, Miami, Tampa, Chicago, Boston, Charlotte, Las Vegas, New York, Cleveland, and Portland. In this case, the covariance matrix is based on five years of monthly returns beginning January 1987 and ending March 2010, and the absorption ratio is based on the first three eigenvectors, roughly $1/5^{\text{th}}$ of the number of assets. The half-time for exponentially-weighted variances is set to two and a half years.

It reveals that the housing market absorption ratio experienced a significant step up from 47.60% in October 1996 to 67.04% in March 1998, just as the national housing bubble got underway. It reached a historic peak of 72.76% in September 1998 and then another peak in July 2004 at 85.63% as the housing bubble continued to inflate. It again reached a historic peak of 89.07% in December 2006 within a few months of the housing bubble peak. Then as the housing bubble burst, the absorption ratio climbed sharply, reaching at an all time high of 94.22% in March 2008. As housing prices stabilized and recovered slightly in 2009, the absorption ratio began to retreat modestly.

It is quite clear from this exhibit that systemic risk in the national housing market increased significantly leading up to the beginning stages of the housing bubble and as the bubble inflated, and that it increased even further after the bubble burst and housing prices tumbled.

The Absorption Ratio and Financial Turbulence

Next we turn to the relationship between systemic risk and financial turbulence. We define financial turbulence as a condition in which asset prices behave in an uncharacteristic fashion given their historical pattern of behavior, including extreme price moves, decoupling of correlated assets, and convergence of uncorrelated assets. We measure financial turbulence as:

$$d_t = (y_t - \mu) \Sigma^{-1} (y_t - \mu)' \quad (2)$$

where,

d_t = turbulence for a particular time period t

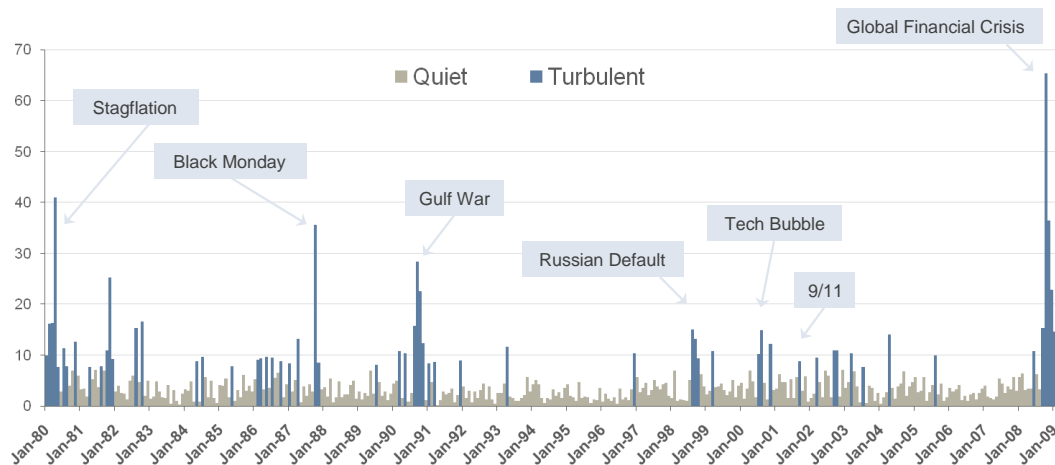
y_t = vector of asset returns for period t

μ = sample average vector of historical returns

Σ = sample covariance matrix of historical returns

Here is how to interpret this formula. By subtracting the historical average from each asset's return we capture the extent to which one or more of the returns was unusually high or low. By multiplying these differences by the inverse of covariance matrix of returns, we both divide by variance, which makes the measure scale independent, and we capture the interaction of the assets. By post multiplying by the transpose of the differences between the asset returns and their averages, we convert this measure from a vector to a single number. Previous research has shown that this statistical characterization of financial turbulence is highly coincident with events in financial history widely regarded as turbulent, as shown in Exhibit 16.¹⁷

Exhibit 16: Financial Turbulence¹⁸

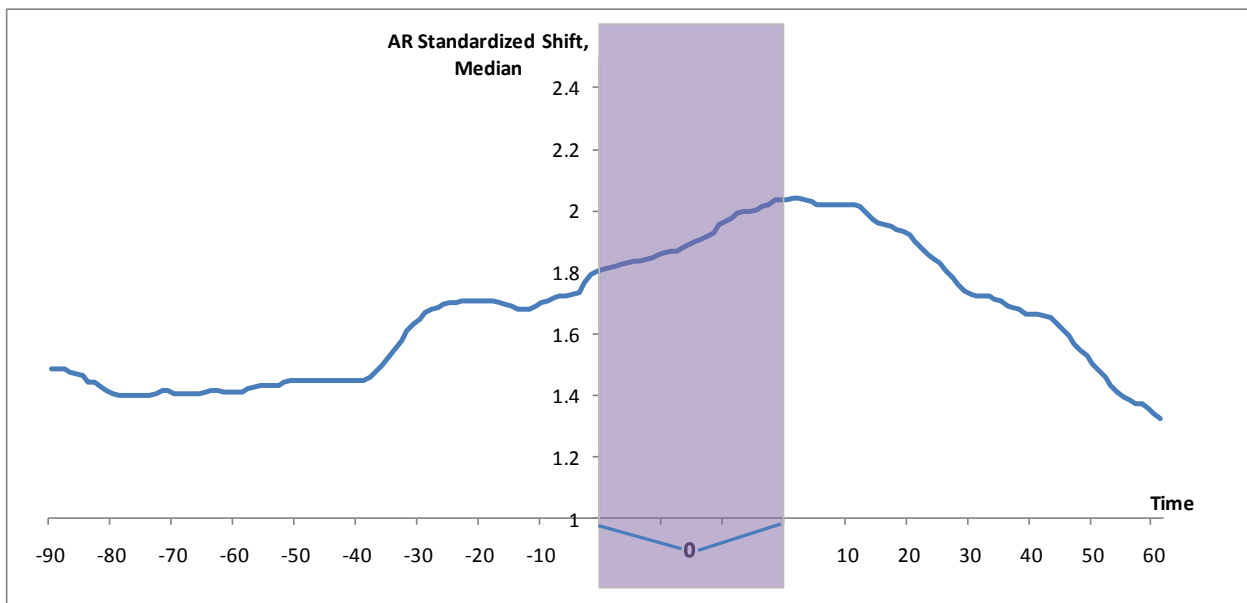


¹⁷ For more about this measure of financial turbulence, including its derivation, empirical properties, and usefulness, see Kritzman, M. and Y. Li, "Skulls, Financial Turbulence, and Risk Management," forthcoming, *Financial Analysts Journal*

¹⁸ This index of financial turbulence is based on daily returns of global stocks, bonds, real estate, and commodities.

In order to measure the connection between systemic risk and financial turbulence, we first identify the 10% most turbulent 30-day periods, based on average daily turbulence, of the MSCI USA stock index covering the period from January 1, 1997 through January 10, 2010. We then synchronize all these turbulent events and observe changes in the 15-day absorption ratio relative to the one-year absorption ratio estimated from industry returns as described earlier, leading up to and following the turbulent events. Exhibit 17 shows the results of this event study.

Exhibit 17: Median Absorption Ratio around Turbulent Periods



Prior to turbulent events in the stock market, the median of the standardized shift in the absorption ratio increased beginning about 40 days in advance of the event, and continued to rise throughout the turbulent periods. It then fell following the conclusion of the turbulent episodes. This evidence suggests that the absorption ratio is an effective precursor of both the inception and conclusion of turbulent episodes, which could prove to be quite valuable. In addition to

persistence, another feature of turbulence is that returns to risk are much lower during turbulent periods than non-turbulent periods.¹⁹

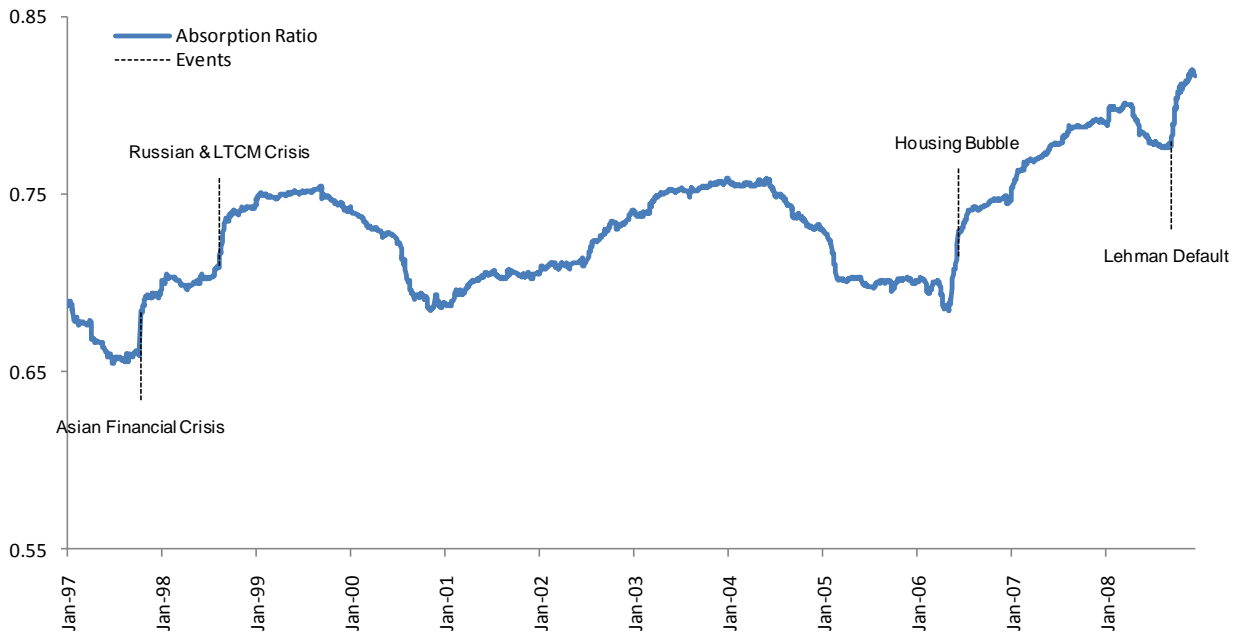
The Absorption Ratio and Global Financial Crises

The previous subsections examined the performance of the absorption ratio in the domestic economy. We now analyze its implications in the global economy. To calculate the global absorption ratio, we collected daily stock market returns for 42 countries (and some regional indexes) from February 1995 to December 2009.

Exhibit 18 shows that the global absorption ratio shifts within a range of 65% to 85% percent. Also, it shows that the global absorption ratio increased in October of 1997 (Hong Kong's speculative attack after the Asian Financial Crises), and August of 1998 (Russian and LTCM collapses); which were two of the most significant emerging market crises in the last 20 years (the other one being the Tequila crisis in 1994 that is outside our sample).

¹⁹ See Kritzman and Li, 2010.

Exhibit 18: Global Absorption Ratio



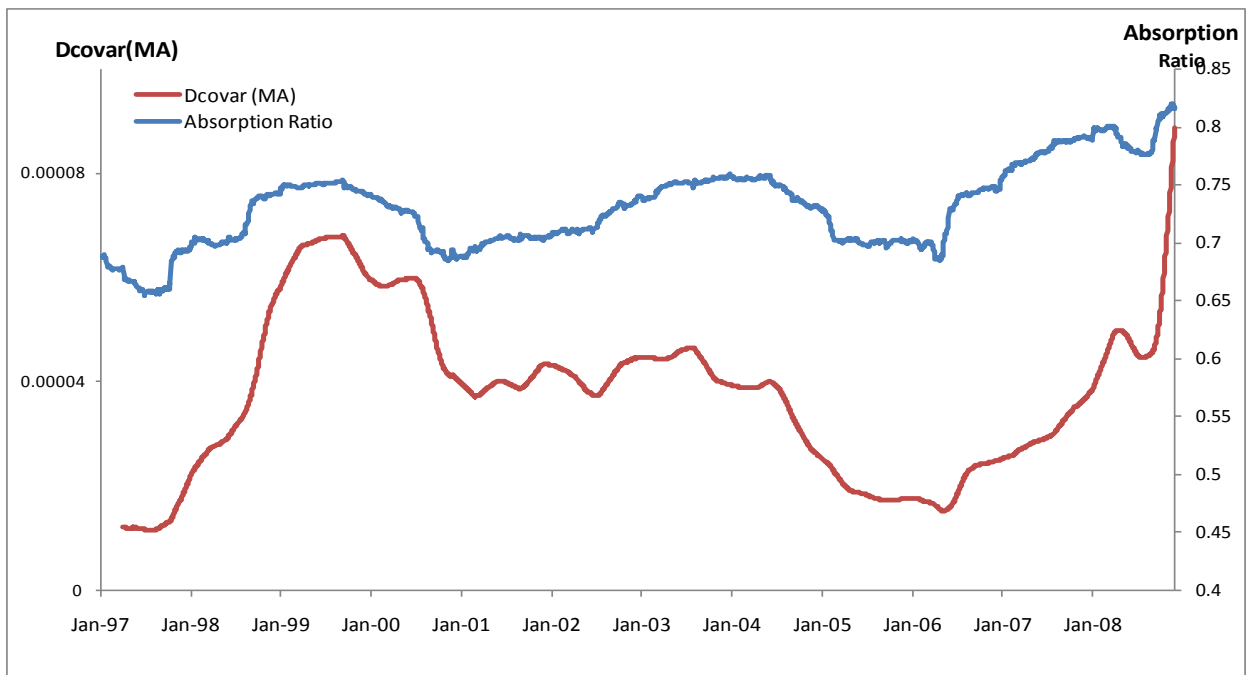
During the recovery in emerging markets (1999-2001), the global absorption ratio decreased. It increased following the boom that came after the severe declines in interest rates that took place after September 11, the accounting standard scandals, and the Dot Com collapse in 2001. Finally, the last significant increase takes place starting in mid 2006 coinciding with the housing bubble, and then the crisis after Lemman’s default. This analysis clearly shows that the global systemic risk of the recent crisis was by far the most severe in recent history

An interesting aspect of the global absorption ratio is that it is highly correlated with structural measures of contagion. In Pavlova and Rigobon (2008), the authors provide a structural model in which asset prices are affected by financial constraints. In that model, the covariance of all asset prices commoves with the degree of the financial constraint. In fact, if all

asset returns are measured on a common currency, then all the covariances move by the exact same amount. This provides a simple indicator of how contagion should take place.

Exhibit 19 shows that the 500 day moving average of the average change in covariance (structural measure of contagion proposed in [Pavlova & Rigobon 2008]) closely coincides with the global absorption ratio. This comparison suggests that international contagion and systemic risk are closely related, which is intuitively pleasing.

Exhibit 19: Global Absorption Ratio and Smoothed Average Change in Covariance



Part IV: Summary

We have introduced a method for inferring systemic risk from asset prices, which we call the absorption ratio. It is equal to the fraction of a set of assets' total variance explained or absorbed by a finite number of eigenvectors. A high absorption ratio implies that financial markets are relatively compact. When markets are compact they are more fragile, because shocks propagate more quickly and broadly. A low absorption ratio suggests that markets are less tightly coupled and therefore less vulnerable to shocks.

Compact markets do not always lead to asset depreciation, but most significant stock market drawdowns were preceded by spikes in the absorption ratio. This suggests that spikes in the absorption are a near necessary but not sufficient condition for market crashes.

We have shown that stock returns are much lower, on average, following spikes in the absorption ratio than they are in the wake of significant declines in the absorption ratio and that investors could have profited by varying exposure to stocks following significant changes in the absorption ratio. We have demonstrated that the absorption ratio of the U.S. housing market provided early signs of the emergence of a national housing bubble, long before the Fed recognized this fact. We have presented evidence showing that increases in the absorption ratio anticipate by more than a month subsequent episodes of financial turbulence. Finally, we have shown that variation in the absorption ratio coincided with many global financial crises and tracked other more complex measures of financial contagion. In short, the absorption ratio appears to serve as an extremely effective measure of systemic risk in financial markets.

Appendix: The Herfindahl Index as a Measure of Implied Systemic Risk

As alternative to our methodology of estimating the absorption ratio as the fraction of total variance explained by a subset of eigenvectors, we considered using the Herfindahl Index to infer systemic risk. Specifically, we squared the fraction of total variance explained by each eigenvector and took the square-root of the sum of these squared values, as shown below.

Insert
$$AR_H = \sqrt{\sum_{i=1}^N \left(\sigma_{Ei}^2 / \sum_{j=1}^N \sigma_{Aj}^2 \right)^2} \quad (3)$$

where,

AR_H : Absorption ratio estimated from Herfindahl index

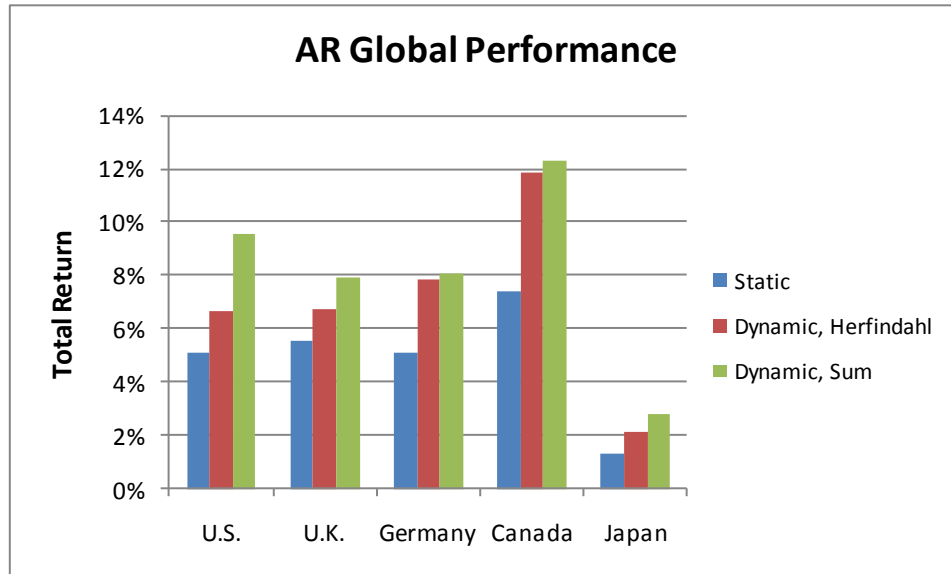
N : number of assets

σ_{Ei}^2 : variance of the i -th eigenvector, sometimes called eigenportfolio

σ_{Aj}^2 : variance of the j -th asset

Exhibit 20 clearly shows that our approach leads to better results than the using a Herfindahl index to estimate systemic risk. Our conjecture is that the less important eigenvectors, which are included in the Herfindahl index, are relatively unstable and introduce noise to the estimate of systemic risk.

Exhibit 20: Herfindahl Index versus Sum of Subset



References

Ang, Andrew and Geert Bekaert. 2002. "International Asset Allocation with Regime Shifts." *Review of Financial Studies*, vol. 15, no. 4 (Fall):1137–1187.

Ang, Andrew, and Joseph Chen. 2002. "Asymmetric Correlations of Equity Portfolios." *Journal of Financial Economics*, vol. 63, no. 3 (March):443–494.

Ang, Andrew, Joseph Chen, and Yuhang Xing. 2002. "Downside Correlation and Expected Stock Returns." Working paper, Columbia University.

Baele, Lieven. 2003. "Volatility Spillover Effects in European Equity Markets: Evidence from a Regime Switching Model." Working paper no. 33, United Nations University, Institute for New Technologies.

Billio, Monica, Getmansky, Mila, Lo, Andrew W. and Pelizzon, Lorian, Measuring Systemic Risk in the Finance and Insurance Sectors (March 10, 2010). MIT Sloan Research Paper No. 4774-10. Available at SSRN: <http://ssrn.com/abstract=1571277>

Butler, Kirt C., and Domingo Castelo Joaquin. 2002. "Are the Gains from International Portfolio Diversification Exaggerated? The Influence of Downside Risk in Bear Markets." *Journal of International Money and Finance*, vol. 21, no. 7 (December):981–1011.

Campbell, Rachel, Kees Koedijk, and Paul Kofman. 2002. "Increased Correlation in Bear Markets." *Financial Analysts Journal*, vol. 58, no. 1 (January/February):87–94.

Cappiello, Lorenzo, Robert F. Engle, and Kevin Sheppard. 2006. "Asymmetric Dynamics in the Correlations of Global Equity and Bond Returns." *Journal of Financial Econometrics*, vol. 4, no. 4 (Fall):385–412

Chow, George, Eric Jacquier, Mark Kritzman, and Kenneth Lowry. 1999. "Optimal Portfolios in Good Times and Bad." *Financial Analysts Journal*, vol. 55, no. 3 (May/June):65–73.

De Bandt, Olivier, and Philipp Hartmann. 2000. "Systemic Risk: A Survey". European Central Bank Working Paper, no. 35.

Ferreira, Miguel A., and Paulo M. Gama. 2004. "Correlation Dynamics of Global Industry Portfolios" Working paper, ISCTE Business School, Lisbon.

Forbes, Kritstin, and Roberto Rigobon. 2006. "No Contagion, only Interdependence: Measuring Stock Market Co-movements"

Goodwin, Thomas H. 1993. "Business-Cycle Analysis With a Markov-Switching Model." *Journal of Business and Economic Statistics*, vol.11, no. 3: 331-339

Gray, Stephen F. 1996. "Modeling the Conditional Distribution of Interest Rates as a Regime-Switching Process." *The Journal of Financial Economics*, vol. 42: 27-62

Hamilton, John. 1989. "A New Approach to the Economic Analysis of Nonstationary Time Series and the Business Cycle." *Econometrica*, vol. 57: 357-384

Hong, Yongmiao, Jun Tu, and Guofu Zhou. 2007. "Asymmetries in Stock Returns: Statistical Tests and Economic Evaluation." *The Review of Financial Studies*, vol. 20, no. 5 (September):1547–1581.

Hyde, Stuart, Don Bredin, and Nghia Nguyen. 2007. "Correlation Dynamics between Asia-Pacific, EU and US Stock Returns." In *Asia-Pacific Financial Market:*

Integration, Innovation and Challenges, *International Finance Review*, vol. 8:39–61.

Hyde, Stuart, Don Bredin, and Nghia Nguyen. 2007. "Correlation Dynamics between Asia-Pacific, EU and US Stock Returns." In *Asia-Pacific Financial Market: Integration, Innovation and Challenges*, International Finance Review, vol. 8:39–61.

International Monetary Fund, "Global Financial Stability Report: Responding to the Financial Crisis and Measuring Systemic Risks", April 2009. www.imf.org

Kim, Chang-Jin. 1993. "Unobserved-Component Time Series Models with Markov-Switching Heteroscedasticity.", *Journal of Business and Economic Statistics*, vol. 11, no. 3: 341-349.

Kritzman, mark and Y. Li. "Skulls, Financial Turbulence, and Risk Management." *Financial Analysts Journal*, forthcoming.

Kritzman, Mark, Kenneth Lowry, and Anne-Sophie Van Royen. 2001. "Risk, Regimes, and Overconfidence." *The Journal of Derivatives*, vol. 8, no. 3 (Spring):32–43.

Kumar, Manmohan S. and Tatsuyoshi Okimoto. 2007. "Dynamics of Persistence in International Inflation Rates." *Journal of Money, Credit and Banking*, vol. 39, no.6: 1458-1479

Lam, Poksank. 2004. "A Markov-Switching Model of GNP Growth with Duration Dependence." *International Economic Review*, vol. 45, no.1: 175-204

Longin, François, and Bruno Solnik. 2001. "Extreme Correlation of International Equity Markets." *The Journal of Finance*, vol. 56, no. 2 (April):649–676.

Luginbuhl, Rob, and Aart de Vos. 1999. "Bayesian Analysis of an Unobserved-Component Time Series Model of GDP With Markov-Switching and Time-Varying Growths." *Journal of Business and Economic Statistics*, vol. 17, no. 4: 456-465.

Pavlova A. and Rigobon R. 2008. "The Role of Portfolio Constraints in the International Propagation of Shocks", *Review of Economic Studies*, 75, pp. 1215-1256.

Smith, Daniel R. 2002. "Markov-Switching and Stochastic Volatility Diffusion Models of Short-Term Interest Rates." *Journal of Business and Economic Statistics*, vol. 20, no. 2: 183-197.

Van Royen, Anne-Sophie. 2002a. "Hedge Fund Index Returns." *Hedge Fund Strategies*, vol. 36 (Fall):111–117.

Van Royen, Anne-Sophie. 2002b. "Financial Contagion and International Portfolio Flows." *Financial analysts Journal*, vol.58, no. 1 (January/February):35–49.