

# Approximating Tracy–Widom distributions

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# Univariate Statistics

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Recall that if  $X_j$  are independent and identically distributed standard normal random variables,  $N(0, 1)$ , then the distribution of

$$\chi_n^2 := X_1^2 + \cdots + X_n^2$$

has density

$$f_n(x) = \begin{cases} \frac{1}{2^{n/2} \Gamma(n/2)} x^{n/2-1} e^{-x/2} & \text{for } x > 0, \\ 0 & \text{for } x \leq 0, \end{cases}$$

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Suppose  $X$  is a  $p \times 1$ -variate normal with  $\mathbb{E}(X) = \mu$  and  $p \times p$  covariance matrix

$\Sigma = \text{cov}(X) := \mathbb{E}((X - \mu) \otimes (X - \mu))$ , denoted  $N_p(\mu, \Sigma)$ .

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If  $\Sigma > 0$  the density function of  $X$  is

$$f_X(x) = (2\pi)^{-p/2} (\det \Sigma)^{-1/2} \exp \left[ -\frac{1}{2} (x - \mu, \Sigma^{-1}(x - \mu)) \right],$$

where  $x \in \mathbb{R}^p$  and  $(\cdot, \cdot)$  is the standard inner product on  $\mathbb{R}^p$ .

# Matrix notation

Introduce a matrix notation: If  $X$  is a  $n \times p$  matrix (the *data matrix*) whose rows  $X_j$  are independent  $N_p(\mu, \Sigma)$  random variables,

$$X = \begin{pmatrix} \leftarrow & X_1 & \rightarrow \\ \leftarrow & X_2 & \rightarrow \\ & \vdots & \\ \leftarrow & X_n & \rightarrow \end{pmatrix},$$



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then we say  $X$  is  $N_p(\mathbf{1} \otimes \mu, I_n \otimes \Sigma)$  where  $\mathbf{1} = (1, 1, \dots, 1)$  and  $I_n$  is the  $n \times n$  identity matrix.

# Multivariate Gamma function

If  $\mathcal{S}_m^+$  is the space of  $p \times p$  positive definite, symmetric matrices, then

$$\Gamma_p(a) := \int_{\mathcal{S}_p^+} e^{-\text{tr}(A)} (\det A)^{a-(p+1)/2} dA$$

where  $\text{Re}(a) > (m-1)/2$  and  $dA$  is the product Lebesgue measure of the  $\frac{1}{2}p(p+1)$  distinct elements of  $A$ .

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$\Gamma_1(a)$  is the usual gamma function  $\Gamma(a)$ .

# Multivariate generalization

**Definition.** If  $A = X^t X$ , where the  $n \times p$  matrix  $X$  is  $N_p(0, I_n \otimes \Sigma)$ , then  $A$  is said to have Wishart distribution with  $n$  degrees of freedom and covariance matrix  $\Sigma$ . We write  $A$  is  $W_p(n, \Sigma)$ .

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**Theorem** (Wishart 1928). If  $A$  is  $W_p(n, \Sigma)$  with  $n \geq p$ , then the density function of  $A$  is

$$\frac{1}{2^{pn/2} \Gamma_p(n/2) (\det \Sigma)^{n/2}} e^{-\frac{1}{2} \text{Tr}(\Sigma^{-1} A)} (\det A)^{(n-p-1)/2}.$$

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For  $p = 1$  and  $\Sigma = 1$  this reduces to the univariate Pearson  $\chi^2$  density. The case  $p = 2$  was obtained by Fisher in 1915 and for general  $p$  by Wishart in 1928 using geometrical arguments. Most modern proofs follow James.

# Importance of Wishart density

Fact: the *sample covariance matrix*,  $S$ , is  $W_p(n, \frac{1}{n}\Sigma)$  where

$$S := \frac{1}{n} \sum_{j=1}^N (X_j - \bar{X}) \otimes (X_j - \bar{X}), \quad N = n + 1,$$

and  $X_j$ ,  $j = 1, \dots, N$ , are independent  $N_p(\mu, \Sigma)$  random vectors, and  $\bar{X} = \frac{1}{N} \sum_j X_j$ .



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Principle component analysis, a multivariate data reduction technique, requires the eigenvalues of the sample covariance matrix; in particular, the largest eigenvalue (the largest principle component variance) is most important.

# Joint pdf for Wishart matrix eigenvalues

**Theorem** (James 1964). *If  $A$  is  $W_p(n, \Sigma)$  with  $n \geq p$  the joint density function of the eigenvalues  $\ell_1, \dots, \ell_p$  of  $A$  is*

$$\frac{\pi^{p^2/2} 2^{-pn/2} (\det \Sigma)^{-n/2}}{\Gamma_p(p/2) \Gamma_p(n/2)} \prod_{j=1}^p \ell_j^{(n-p-1)/2} \prod_{j < k} (\ell_j - \ell_k) \\ \cdot \int_{\mathbb{O}(p)} e^{-\frac{1}{2} \text{Tr}(\Sigma^{-1} H L H^t)} dH$$

*where  $\mathbb{O}(p)$  is the orthogonal group of  $p \times p$  matrices,  $dH$  is normalized Haar measure and  $L$  is the diagonal matrix  $\text{diag}(\ell_1, \dots, \ell_p)$ . (We take  $\ell_1 > \ell_2 > \dots > \ell_p$ .)*

# Evaluation of joint pdf

$$\frac{\pi^{p^2/2} 2^{-pn/2} (\det \Sigma)^{-n/2}}{\Gamma_p(p/2) \Gamma_p(n/2)} \prod_{j=1}^p \ell_j^{(n-p-1)/2} \prod_{j < k} (\ell_j - \ell_k) \\ \cdot \int_{\mathbb{O}(p)} e^{-\frac{1}{2} \text{Tr}(\Sigma^{-1} H L H^t)} dH$$

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No known closed formula though James and Constantine developed the theory of *zonal polynomials* which allow one to write infinite series expansions for this integral.

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For complex Wishart matrices, the group integral is over the unitary group  $\mathbb{U}(p)$ ; this integral can be evaluated using the Harish-Chandra-Itzykson-Zuber integral (Zinn–Justin '03).

# Evaluation of joint pdf 3

There is one important case where the integral can be (trivially) evaluated.

**Corollary.** *If  $\Sigma = I_p$ , then the joint density simplifies to*

$$\frac{\pi^{p^2/2} 2^{-pn/2} (\det \Sigma)^{-n/2}}{\Gamma_p(p/2) \Gamma_p(n/2)} \prod_{j=1}^p \ell_j^{(n-p-1)/2} \exp \left( -\frac{1}{2} \sum_j \ell_j \right) \prod_{j < k} (\ell_j - \ell_k)$$



# Connection to RMT

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**Definition.**

$$\mu_{np} = \left( \sqrt{n-1} + \sqrt{p} \right)^2 ,$$

$$\sigma_{np} = \left( \sqrt{n-1} + \sqrt{p} \right) \left( \frac{1}{\sqrt{n-1}} + \frac{1}{\sqrt{p}} \right)^{1/3} .$$

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**Theorem** (Johnstone '01). *Under the null hypothesis  $\Sigma = I_p$ , if  $n, p \rightarrow \infty$  such that  $n/p \rightarrow \gamma, 0 < \gamma < \infty$ , then*

$$\frac{\ell_1 - \mu_{np}}{\sigma_{np}} \xrightarrow{\mathcal{D}} F_1(s, 1).$$

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**Theorem** (Soshnikov, '02). *If  $n, p \rightarrow \infty$  such that  $n/p \rightarrow \gamma, 0 < \gamma < \infty$ , then*

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Subsequently, El Karoui ('03) extended Soshnikov's Theorem to  $0 < \gamma \leq \infty$ ; extension to  $\gamma = \infty$  is important for modern statistics where  $p \gg n$  arises in applications.

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Redefine the  $n \times p$  matrices  $X = \{x_{i,j}\}$  such that  $A = X^t X$  to satisfy

1.  $\mathbb{E}(x_{ij}) = 0, \mathbb{E}(x_{ij}^2) = 1.$
2. The random variables  $x_{ij}$  have symmetric laws of distribution.
3. All even moments of  $x_{ij}$  are finite, and they decay at least as fast as a Gaussian at infinity:  
$$\mathbb{E}(x_{ij}^{2m}) \leq (\text{const } m)^m.$$
4.  $n - p = O(p^{1/3}).$

# Connection to RMT 4

With these assumptions,

**Theorem** (Soshnikov '02).

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Deift and Gioev ('05), building on the work of Widom ('99), proved  $F_1$  universality when the Gaussian weight function  $\exp(-x^2)$  is replaced by  $\exp(-V(x))$  where  $V$  is an even degree polynomial with positive leading coefficient.

# Fredholm det. representation ( $\beta = 2$ )

In the unitary case ( $\beta = 2$ ), define the trace class operator  $K_2$  on  $L^2(s, \infty)$  with *Airy kernel*

$$K_{\text{Ai}}(x, y) := \frac{\text{Ai}(x) \text{Ai}'(y) - \text{Ai}'(x) \text{Ai}(y)}{x - y} = \int_0^\infty \text{Ai}(x+z) \text{Ai}(y+z) dz$$

and associated Fredholm determinant,  $0 \leq \lambda \leq 1$ ,

$$D_2(s, \lambda) = \det(I - \lambda K_2).$$

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Then

$$F_2(s, m+1) - F_2(s, m) = \frac{(-1)^m}{m!} \frac{d^m}{d\lambda^m} D_2(s, \lambda) \Big|_{\lambda=1}, \quad m \geq 0,$$

where  $F_2(s, 0) \equiv 0$

# Fredholm det. representation ( $\beta = 1, 4$ )

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$$S_4(x, y) = K_{\text{Ai}}(x, y) - \frac{1}{2} \text{Ai}(x, y) \int_y^\infty \text{Ai}(z) dz,$$

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$$\varepsilon(x - y) = \frac{1}{2} \operatorname{sgn}(x - y).$$

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Recall that in the unitary ( $\beta = 2$ ) case we have the recurrence

$$F_2(s, m+1) - F_2(s, m) = \frac{(-1)^m}{m!} \frac{d^m}{d\lambda^m} D_2(s, \lambda) \Big|_{\lambda=1}, \quad m \geq 0,$$

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Similarly in the orthogonal and symplectic ( $\beta = 1, 4$ ) case

$$F_\beta(s, m+1) - F_\beta(s, m) = \frac{(-1)^m}{m!} \frac{d^m}{d\lambda^m} D_\beta^{1/2}(s, \lambda) \Big|_{\lambda=1}, \quad m \geq 0,$$

where  $\beta = 1, 4$  and  $F_\beta(s, 0) \equiv 0$

# Painlevé representations

**Theorem** (Clarkson, McLeod, '88). : *There exist a unique solution  $q(x, \lambda)$  to the Painlevé II equation*

$$q'' = x q + 2 q^3$$

*such that  $q \rightarrow \sqrt{\lambda} \text{Ai}$  as  $x \rightarrow \infty$  and  $\text{Ai}(x)$  is the solution to the Airy equation that decays like  $\frac{1}{2}\pi^{-1/2}x^{-1/4} \exp\left(-\frac{2}{3}x^{3/2}\right)$  at  $+\infty$ .*

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**Definition.**

$$\mu(s, \lambda) := \int_s^\infty q(x, \lambda) dx,$$

$$\tilde{\lambda} := 2\lambda - \lambda^2,$$



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**Theorem** (Dieng '05).

$$D_1(s, \lambda) = D_2(s, \tilde{\lambda}) \frac{\lambda - 1 - \cosh \mu(s, \tilde{\lambda}) + \sqrt{\tilde{\lambda}} \sinh \mu(s, \tilde{\lambda})}{\lambda - 2} ,$$

$$D_4(s, \lambda) = D_2(s, \lambda) \cosh^2 \left( \frac{\mu(s, \lambda)}{2} \right) .$$

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**Theorem** (Dieng '05).

$$D_1(s, \lambda) = D_2(s, \tilde{\lambda}) \frac{\lambda - 1 - \cosh \mu(s, \tilde{\lambda}) + \sqrt{\tilde{\lambda}} \sinh \mu(s, \tilde{\lambda})}{\lambda - 2},$$

$$D_4(s, \lambda) = D_2(s, \lambda) \cosh^2 \left( \frac{\mu(s, \lambda)}{2} \right).$$

**Software:** <http://math.arizona.edu/~momar>

# Edgeworth expansions

If  $S_n$  is a sum of i.i.d. random variables  $X_j$ , each with mean  $\mu$  and variance  $\sigma^2$ , the distribution  $F_n$  of the normalized random variable  $(S_n - n\mu)/(\sigma\sqrt{n})$  satisfies the Edgeworth expansion

$$F_n(x) - \Phi(x) = \phi(x) \sum_{j=3}^r n^{-\frac{1}{2}j+1} R_j(x) + o(n^{-\frac{1}{2}r+1})$$

uniformly in  $x$ ;  $\Phi$  is the standard normal distribution with density  $\phi$ , and  $R_j$  are polynomials depending only on  $\mathbb{E}(X_j^k)$  but not on  $n$  and  $r$  (or the underlying distribution of the  $X_j$ ).

# Edgeworth expansions 2

Following Tracy and Widom we define

$$u_i \quad := \quad u_i(s) = \int_s^\infty q(x) x^i \operatorname{Ai}(x) dx,$$

$$v_i \quad := \quad v_i(s) = \int_s^\infty q(x) x^i \operatorname{Ai}'(x) dx$$

and

$$w_i := w_i(s) = \int_s^\infty q'(x) x^i \operatorname{Ai}'(x) dx + u_0(s) v_i(s)$$

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Let  $c$  be an arbitrary constant (tuning parameter) and

$$E(s) = 2w_1 - 3u_2 + (-20c^2 + 3)v_0 + u_1v_0 - u_0v_1 + u_0v_0^2 - u_0^2w_0,$$

# Edgeworth expansions 3

**Theorem** (Choup '06). *Setting*

$$t = (2(n + c))^{\frac{1}{2}} + 2^{-\frac{1}{2}} n^{-\frac{1}{6}} s$$

*Then as  $n \rightarrow \infty$*

$$F_{n,2}(t) = F_2(s) \left\{ 1 + c u_0(s) n^{-\frac{1}{3}} - \frac{1}{20} E(s) n^{-\frac{2}{3}} \right\} + O(n^{-1})$$

*uniformly in  $s$ .*



# Edgeworth expansions 4

One key consequence of Choup's work is the expansion for  $R_n(x, y) = (I - K_{n,2}\chi_{(s,\infty)})$ .

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$$Q_i(s) := (I - K_{\text{Ai}})X^i \text{Ai}(X) \quad \text{and} \quad P_i(s) := (I - K_{\text{Ai}})X^i \text{Ai}'(X)$$

# Edgeworth expansions 4

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Then

$$\begin{aligned} R_n(x, y) = & R(X, Y) - c Q \otimes Q n^{-1/3} + \\ & + \frac{n^{-2/3}}{20} [P_1 \otimes P + P \otimes P_1 - Q_2 \otimes Q - Q_1 \otimes Q_1 \\ & - Q \otimes Q_2 + \frac{3 - 20 c^2}{2} (P \otimes Q + Q \otimes P) + \\ & + 20 c^2 u_0(s) Q \otimes Q] + O(n^{-1}) \end{aligned}$$

# Edgeworth expansions 5

In the  $\beta = 1, 4$  cases we have the formulas (Tracy and Widom)

$$\begin{aligned} F_{n,1}(t) &= (1 - \tilde{v}_\varepsilon) \left(1 - \frac{1}{2} \mathcal{R}_1\right) - \frac{1}{2} (q_\varepsilon - c_\varphi) \mathcal{P}_1, \\ F_{n,4}(t/\sqrt{2}) &= (1 - \tilde{v}_\varepsilon) \left(1 + \frac{1}{2} \mathcal{R}_4\right) + \frac{1}{2} q_\varepsilon \mathcal{P}_4, \end{aligned}$$

# Edgeworth expansions 5

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All quantities are expressible in terms of  $R_n(x, y)$  and other quantities whose known expansions can be used in the above formulas. Details are rather messy, so we leave them for private discussions.