

Improved Concentration Bounds for Count-Sketch

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Count-Sketch: a classic streaming algorithm

Charikar, Chen, Farach-Colton 2002

- Solves “heavy hitters” problem
- Estimate a vector $x \in \mathbb{R}^n$ from low dimensional sketch $Ax \in \mathbb{R}^m$.
- Nice algorithm
 - ▶ Simple
 - ▶ Used in Google’s MapReduce standard library
- [CCF02] bounds the *maximum* error over all coordinates.
- We show, for the same algorithm,
 - ▶ *Most* coordinates have asymptotically better estimation accuracy.
 - ▶ The *average* accuracy over many coordinates will be asymptotically better *with high probability*.
 - ▶ Experiments show our asymptotics are correct.
- Caveat: we assume fully independent hash functions.

Outline

1 Robust Estimation of Symmetric Variables

- Lemma
- Relevance to Count-Sketch

2 Electoral Colleges and Direct Elections

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3 Experiments!

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Estimating a symmetric random variable's mean

\mathcal{X}

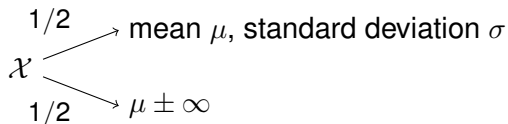
- Unknown distribution \mathcal{X} over \mathbb{R} , symmetric about unknown μ .
 - ▶ Given samples $x_1, \dots, x_R \sim \mathcal{X}$.
 - ▶ How to estimate μ ?

Estimating a symmetric random variable's mean

\mathcal{X} \nearrow mean μ , standard deviation σ

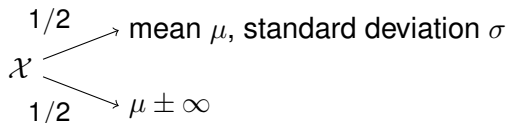
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- Mean:
 - ▶ Converges to μ as σ/\sqrt{R} .

Estimating a symmetric random variable's mean



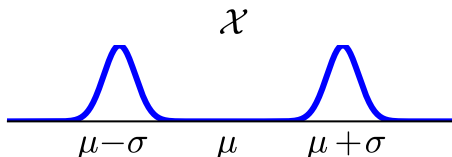
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Estimating a symmetric random variable's mean



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 - ▶ Given samples $x_1, \dots, x_R \sim \mathcal{X}$.
 - ▶ How to estimate μ ?
- Mean:
 - ▶ Converges to μ as σ/\sqrt{R} .
 - ▶ No robustness to outliers
- Median:
 - ▶ Extremely robust
 - ▶ Doesn't necessarily converge to μ .

Estimating a symmetric random variable's mean



- Median doesn't converge
- Consider: median of pairwise means

$$\hat{\mu} = \operatorname{median}_{i \in \{1, 3, 5, \dots\}} \frac{X_i + X_{i+1}}{2}$$

- ▶ Converges as $O(\sigma/\sqrt{R})$, even with outliers.
- That is: median of $(\mathcal{X} + \mathcal{X})$ converges.

[See also: Hodges-Lehmann estimator.]

Why does median converge for $\mathcal{X} + \mathcal{X}$?

- WLOG $\mu = 0$.
- Define the Fourier transform $\mathcal{F}_{\mathcal{X}}$ of \mathcal{X} :

$$\mathcal{F}_{\mathcal{X}}(t) = \mathbb{E}_{x \sim \mathcal{X}} [\cos(\overset{2\pi \approx 6.28}{\tau} xt)]$$

(standard Fourier transform of PDF, specialized to symmetric \mathcal{X} .)

- Convolution \iff multiplication
 - ▶ $\mathcal{F}_{\mathcal{X}+\mathcal{X}}(t) = (\mathcal{F}_{\mathcal{X}}(t))^2 \geq 0$ for all t .

Theorem

Let \mathcal{Y} be symmetric about 0 with $\mathcal{F}_{\mathcal{Y}}(t) \geq 0$ for all t and $\mathbb{E}[Y^2] = \sigma^2$.
Then for all $\epsilon \leq 1$,

$$\Pr[|y| \leq \epsilon\sigma] \gtrsim \epsilon$$

Standard Chernoff bounds: median y_1, \dots, y_R converges as σ/\sqrt{R} .

Proof

Theorem

Let $\mathcal{F}_y(t) \geq 0$ for all t and $\mathbb{E}[Y^2] = 1$. Then for all $\epsilon \leq 1$,

$$\Pr[|y| \leq \epsilon] \gtrsim \epsilon.$$

$$\mathcal{F}_y(t) = \mathbb{E}[\cos(\tau yt)] \geq 1 - \frac{\tau^2}{2} t^2$$

$$\Pr[|y| \leq \epsilon] = \mathcal{Y} \cdot \underbrace{\hspace{1.5cm}}_{\epsilon} \Bigg\} 1$$

$$\geq \mathcal{Y} \cdot \underbrace{\hspace{1.5cm}}_{\epsilon} \Bigg\} 1$$

$$= \mathcal{F}_y \cdot \underbrace{\hspace{1.5cm}}_{1/\epsilon} \Bigg\} \epsilon$$

$$\geq \underbrace{\hspace{1.5cm}}_{0.2} \Bigg\} 1 \cdot \underbrace{\hspace{1.5cm}}_{1/\epsilon} \Bigg\} \epsilon \gtrsim \epsilon.$$

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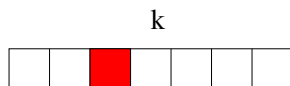
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3 Experiments!

Count-Sketch

- Want to estimate $x \in \mathbb{R}^n$ from small “sketch.”
- Hash to k buckets and sum up with random signs
Choose random $h : [n] \rightarrow [k], s : [n] \rightarrow \{\pm 1\}$. Store

$$y_j = \sum_{i: h(i)=j} s(i)x_i$$



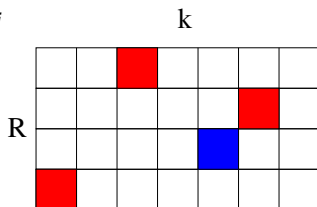
- Can estimate x_i by $\tilde{x}_i = y_{h(i)}s(i)$.

Count-Sketch

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Choose random $h : [n] \rightarrow [k], s : [n] \rightarrow \{\pm 1\}$. Store

$$y_j = \sum_{i: h(i)=j} s(i)x_i$$

- Can estimate x_i by $\tilde{x}_i = y_{h(i)}s(i)$.
- Repeat R times, take the median.
- For each row,



$$\tilde{x}_i - x_i = \sum_{j \neq i} \begin{cases} \pm x_j & \text{with probability } 1/k \\ 0 & \text{otherwise} \end{cases}$$

- Symmetric, non-negative Fourier transform.

Count-Sketch Analysis

Let

$$\sigma^2 = \frac{1}{k} \min_{k\text{-sparse } x_{[k]}} \|x - x_{[k]}\|_2^2$$

be the “typical” error for a single row of Count-Sketch with k columns.

Theorem

For the any coordinate i , we have for all $t \leq R$ that

$$\Pr[|\hat{x}_i - x_i| > \sqrt{\frac{t}{R}}\sigma] \leq e^{-\Omega(t)}.$$

(CCF02: $t = R = O(\log n)$ case; $\|\hat{x} - x\|_\infty \lesssim \sigma$ w.h.p.)

Corollary

Excluding $e^{-\Omega(R)}$ probability events, we have for each i that

$$\mathbb{E}[(\hat{x}_i - x_i)^2] = \sigma^2/R$$

Estimation of *multiple* coordinates?

- What about the average error on a *set* S of k coordinates?
- Linearity of expectation: $\mathbb{E}[\|\hat{x}_S - x_S\|_2^2] = \frac{O(1)}{R} k \sigma^2$.
- Does it concentrate?

$$\Pr[\|\hat{x}_S - x_S\|_2^2 > \frac{O(1)}{R} k \sigma^2] < p = ???$$

- By expectation: $p = \Theta(1)$.
- If independent: $p = e^{-\Omega(k)}$.
- Sum of many variables, but not independent...
- Chebyshev's inequality, bounding covariance of error:
 - ▶ Feasible to analyze (though kind of nasty).
 - ▶ Ideally get: $p = 1/\sqrt{k}$.
 - ▶ We can get $p = 1/k^{1/14}$.
- Can we at least get “high probability,” i.e. $1/k^c$ for arbitrary constant c ?

Boosting the error probability

in a black box manner

- We know that $\|\widehat{x}_S - x_S\|_2$ is “small” with all but $k^{-1/14}$ probability.
- Way to get all but k^{-c} probability: repeat $100c$ times and take the median of results.
 - ▶ With all but k^{-c} probability, $> 75c$ of the $\widehat{x}_S^{(i)}$ will have “small” error.
 - ▶ Median of results has at most $3 \times$ “small” total error.
- But resulting algorithm is stupid:
 - ▶ Run count-sketch with $R' = O(cR)$.
 - ▶ Arbitrarily partition into blocks of R rows.
 - ▶ Estimate is median (over blocks) of median (within block) of individual estimates.
- Can we show that the direct median is as good as the median-of-medians?

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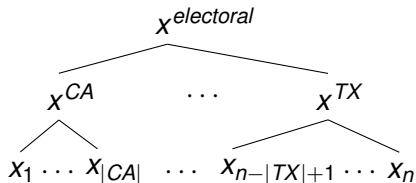
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Electoral Colleges

- Suppose you have a two-party election for k offices.
 - ▶ Voters come from a distribution \mathcal{X} over $\{0, 1\}^k$.
 - ▶ “True” majority slate of candidates $\bar{x} \in \{0, 1\}^k$.
 - ▶ Election day, receive ballots $x_1, \dots, x_n \sim \mathcal{X}$.
- How to best estimate \bar{x} ? For each office,



- Is x^{majority} better than $x^{\text{electoral}}$ in every way? Is

$$\Pr[\|x^{\text{majority}} - \bar{x}\| > \alpha] \leq \Pr[\|x^{\text{electoral}} - \bar{x}\| > \alpha]$$

for all $\alpha, \|\cdot\|$?

Electoral Colleges

- Is x^{majority} better than $x^{\text{electoral}}$ in every way, so

$$\Pr[\|x^{\text{majority}} - \bar{x}\| > \alpha] \leq \Pr[\|x^{\text{electoral}} - \bar{x}\| > \alpha]$$

for all α , $\|\cdot\|$?

- Don't know, but

Theorem

$$\Pr[\|x^{\text{majority}} - \bar{x}\| > 3\alpha] \leq 4 \cdot \Pr[\|x^{\text{electoral}} - \bar{x}\| > \alpha]$$

for all p -norms $\|\cdot\|$.

Proof

Theorem

$$\Pr[\|x^{\text{majority}} - \bar{x}\| > 3\alpha] \leq 4 \cdot \Pr[\|x^{\text{electoral}} - \bar{x}\| > \alpha]$$

for all p -norms $\|\cdot\|$.

Follows easily from:

Lemma (median³)

For any $x_1, \dots, x_n \in \mathbb{R}^k$, we have

$$\underset{\text{partitions into states}}{\text{median}} \quad \underset{\text{states}}{\text{median}} \quad \underset{\text{within state}}{\text{median}} \quad x_i = \underset{\text{populace}}{\text{median}} \quad x_i$$

(With $4p$ failure probability, $3/4$ of partitions have error at most α ; then their median has error 3α .)

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Concentration for sets

- We know that a “median-of-medians” variant of Count-Sketch would give good estimation of sets with high probability.
- Therefore the standard Count-Sketch would as well.

Theorem

For any constant c , we have for any set S of coordinates that

$$\Pr[\|\hat{x}_S - x_S\|_2 > O(\sqrt{\frac{|S|}{R}}\sigma)] \lesssim |S|^{-c}.$$

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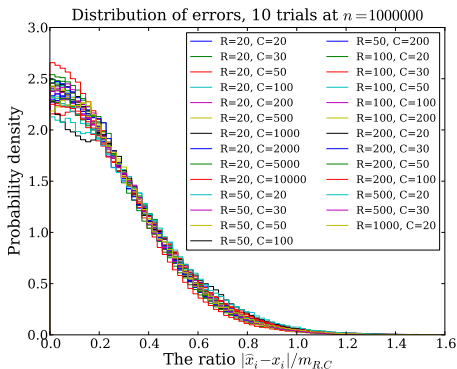
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Experiments

- Claims
 - 1 Individual coordinates have error that concentrates like a Gaussian with standard deviation σ/\sqrt{R} .
 - 2 Sets of coordinates have error $O(\sigma\sqrt{k/R})$ with high probability.
- Evaluate on power-law distribution with typical parameters.

Experiments

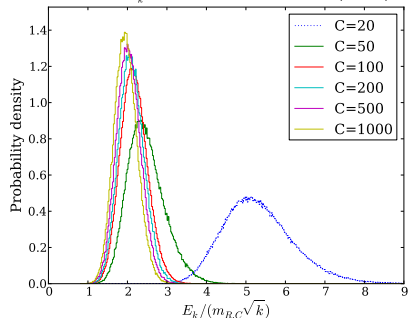
- 1 Individual coordinates have error that concentrates like a Gaussian with standard deviation σ/\sqrt{R} .
- Compare observed error to expected error for various R, C .



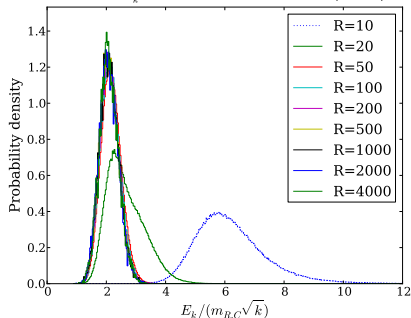
Experiments

- 2 Sets of coordinates have error $O(\sigma\sqrt{k/R})$ with high probability. (for large enough R, C)
- Compare observed error to expected error for various R, C .

Distribution of E_k for various C with $n=10000, k=25, R=50$



Distribution of E_k for various R with $n=10000, k=25, C=100$



Conclusions

- We present an improved analysis of Count-Sketch, a classic algorithm used in practice.
- Experiments show it gives the right asymptotics
- More applications of our lemmas?
- Independence?

Thank You

